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# Ex Post Portfolio Performance with Predictable Skewness and Kurtosis 

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# Ex Post Portfolio Performance with Predictable Skewness and Kurtosis* 

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#### Abstract

This paper examines the ex-post performance of optimal portfolios with predictable returns, when the investor horizon ranges from one month to ten years. Due to the investor's ability to anticipate shifts from bull to bear markets, predictability involves the risk premium, volatility and correlations, and may extend to third and fourth moments. We analyze three different equity portfolios datasets, each covering more than eight indexes, including the commonly used US Industry and International Book-toMarket portfolios. Allowing for regimes improves portfolio performance for at least a subset of investment horizons in all datasets. Despite large non-normalities in both the Industry and the BM dataset, gains from predicting higher order moments obtain only in the latter - where third rather than fourth moments matter.

The equally weighted strategy usually leads to lower ex-post performance measures than optimizing ones, despite simple econometrics and power utility preferences underlying optimal strategies.


Key words: Stock Market Regimes, Return Predictability, Skew and Kurtosis, Equity Diversification JEL code: G11, F37, C22, C51.

## 1. Introduction

Risk-adjusted profits of portfolio managers derive from their ability in forecasting returns out-of-sample. Recently, Bossaerts and Hillion (1999), Ang and Bekaert (2007) and Goyal and Welch (2008) cast doubts on prevailing linear methods for predicting out-of-sample, which are reinforced by the inability of optimizing strategies in obtaining out-of-sample gains relative to a naive equally-weighted strategy (DeMiguel, Garlappi

[^0]and Uppal, 2009). Avramov and Chordia (2006) do find large out-of-sample gains ${ }^{1}$. However they consider individual stocks as opposed to diversified equity portfolios, which favour optimal strategies over simpler ones.

Importantly, all these papers restrict attention to mean-variance preferences in computing out-of-sample welfare gains, thus overlooking the fact that investors appear to care about both asymmetries and tails of the wealth distribution - as indicated by the asset pricing literature (Harvey and Siddique, 2000; Dittmar, 2002; Kumar, 2005; Guidolin and Timmermann, 2008a). This leaves open the possibility of out-of-sample welfare gains deriving from the impact of predictable higher order moments of stock returns on optimal portfolio composition.

Moreover, there is mounting evidence that non-linear models in general and Markov-switching models in particular provide superior fit. Ang and Chen (2002) report that regime switching models replicate the asymmetries in correlations observed in US stock returns better than GARCH-M. Lettau and Van Nieuwerburgh (2008) suggest that the presence of changing steady state means in the dividend-price ratio may explain why it proves so difficult to predict stock returns out of sample with such ratio. Guidolin and Nicodano (2008) find that, both in-sample and out-of-sample, regime switching models with time-varying covariance matrix fare as well as or better than multivariate ARCH models in an international dataset of size indexes.

This paper provides extensive evidence of out-of-sample performance of optimal portfolio strategies in three datasets that are commonly used by both academics and the industry, at monthly frequencies. We analyze ten US industry portfolios (IND) and eleven Book-to-Market (BM) international portfolios, along with eight international and emerging (IE) market stock indexes. In most cases we find out-of-sample gains relative to the equally-weighted strategies for investors who have one period horizons, as those studied by DeMiguel et al.(2009).

Several papers already indicate that predicting higher order moments changes the composition of optimal portfolios (e.g. Ang and Bekaert (2002), Guidolin and Timmermann (2008a), Guidolin and Nicodano (2008), Jondeau and Rockinger (2009)), because investors overweight equity indexes that increase positive wealth skewness and reduce excess wealth kurtosis relative to mean-variance portfolios. These papers - with the exception of Ang and Bekaert (2002) and Guidolin and Timmermann (2008a) - use however weekly data, which amplify the importance of higher order moments relative the commonly used monthly returns. Moreover, each focusses on one set of equity indexes only. Importantly, they offer little discussion of out-ofsample gains and - when they do (as in Jondeau and Rockinger (2009)) - they do not explore the effects on predictability on longer run portfolio performance. ${ }^{2}$ A second contribution of our paper consists in analyzing how the investor horizon, which ranges from 1 month to 10 years, affects ex-post gains.

The prevailing linear forecasting methods - such as those in DeMiguel et al. (2009) - describe stock returns as randomly fluctuating around one mean return with one given volatility. Our portfolio strategies are instead based on models for returns that allow stock markets to persistently remain in either a bear or a bull regime. If the US stock market is in a bear regime, future returns of a given equity portfolio will be

[^1]expected to fluctuate around a given mean return with a given volatility - unless the stock market moves to a bull regime, which may happen with a positive probability that is updated on the basis of upcoming information. In that case, future returns fluctuate around a higher mean return with a lower volatility unless the stock market returns to the bear regime. In our analysis, this representation fits the return data better than the gaussian i.i.d. representation according to standard statistical tests for both US Industry and International BM data. ${ }^{3}$

A state dependent return representation also has a number of advantages from the point of view of portfolio management. Given that returns are assumed to be normal conditional on a given regime, assets are characterized by the familiar expected return - variance statistics in each regime. It is therefore immediate to identify a defensive industry as one having a relatively high return in the bear state, compared to other industries. It also allows to generalize the concept to higher order moments. A truly defensive industry also contributes to increase the skewness of wealth, i.e. has a relatively low variance in the bear regime, and to reduce wealth kurtosis by displaying relatively low variance in highly volatile bear markets. North American stocks and Energy appear to be truly defensive portfolios in our data sets.

There are other well known econometric advantages from using such regime-switching representation. First, the data identify the number of stock market regimes, without the econometrician having to impose them exogenously. Second, it is possible to estimate higher order moments more precisely with a limited amount of observations, because they are a function of the transition probabilities plus conditional means and covariance terms (Timmermann, 2000). Thus, measuring the skewness requires the estimation of fewer parameters than a traditional representation. Third, it is possible to nest other simpler forecasting models as special cases of a general Markov-switching process.

Last but not least, methods that account for systematic skewness and kurtosis in a regime-switching setting are often cumbersome and/or do not allow for consideration of several securities ${ }^{4}$ : this may prevent their use by investors. Our paper uses a tractable approach developed by Guidolin and Timmermann (2008b) which is convenient to implement in the presence of non-normalities and large asset menus.

Another related literature deals with predicting and timing volatility on daily data (see Fleming et al., 2001, and references therein), assuming constant expected returns given the short horizon under scrutiny. Here we investigate whether there are economic gains from predicting and timing up to the fourth moment, all of which are likely to vary over a monthly - or longer - horizon.

The rest of the paper proceeds as follows. We describe the optimal asset allocation problem in Section 2. Section 2.1. contains a description of the portfolio strategies under analysis, whereas the remaining ones provide technical details on the solution method as well as on the specification test concerning the return generating process. Section 3 describes our data set and the distribution of returns conditional on two states

[^2]of the market. Section 4 reports results concerning the composition of optimal portfolios and their ex-post performance.

## 2. Optimal Asset Allocation

### 2.1. Portfolio Strategies and Overview

We summarize the portfolio strategies under scrutiny in Table 0 . We allow the investor to have alternative preferences over moments of the distribution of final wealth. MV denotes mean-variance preferences, MVS indicates an investor who also likes positive skewness of wealth while a MVSK investor also dislikes fat tails in the wealth distribution.

Returns on stock indexes are allowed to follow two alternative processes. One is the traditional GaussianI.I.D. case, which we indicate with $k=1$. Under such distribution, differences in optimal portfolio strategies across investors disappear, because both the skewness and the excess kurtosis of wealth are zero. However, such return distribution is rejected by both Industry and Book-to-Market data when we perform specification tests (see Table 2). Data support an alternative generating process, which we label $k=2$, indicating that returns are normal conditional on two states of the market. In one such state, that we name "bull", expected returns are higher and volatilities are lower than in the "bear" state (see Table 3). Furthermore, it is possible to measure the contribution of each stock index not only to the variance of wealth, but to its skewness and its kurtosis as well. To this end, Table 4 reports the co-skewness and co-kurtosis matrices alongside the conventional covariance measures.

We combine alternative preferences with these two return distributions, as displayed in Table 1. And, for each combination, we compute optimal portfolios (Tables 5) and ex-post performance (Tables 6). When regimes are allowed for, optimal portfolios can be studied along different dimensions. First, we have the "average" portfolio composition when the investor does not know which state the market is in, but attributes to each its long-run probability. Then we have portfolio composition when the market is bear and when it is bull. Finally, we can measure how each portfolio share changes as the probability of being in a bear/bull state is updated by the investor (Figures 2 and 3 ).

We measure ex-post performance through three different indicators. One is the expected return to volatility ratio (Sharpe ratio), which is insensitive to both skewness and kurtosis of wealth. Thus, a portfolio strategy that increases downside risk is given the same grade as another one with the same Sharpe ratio that does not. This does not happen with the Sortino ratio, which falls when downside risk increases. The third performance metric is the certainty equivalent of maximum utility, which - in the case of a MVSK investor - also captures the higher moments of wealth.

Thus, we are able to assess whether the Sharpe/Sortino ratios of a MVSK investor exceeds the one of a MV investor. We can also analyze whether a MV investor would be better off considering time-varying mean and variances, by accounting for regimes. And we quantify the costs for a MVSK investor of using a Gaussian return distribution instead of the regime-switching one.

The following subsections provide technical details, and the reader who is interested only in empirical results may skip it.

### 2.2. Investor Preferences

This section describes the investor's objectives and the return generating process and goes on to characterize the method used to solve for the optimal asset allocation. We are interested in studying the asset allocation problem at time $t$ for an investor with a $T$-period investment horizon. Suppose that the investor's utility function $U\left(W_{t+T} ; \boldsymbol{\theta}\right)$ only depends on wealth at time $t+T, W_{t+T}$, and its shape is captured through the parameters in $\boldsymbol{\theta}$. The investor maximizes expected utility by choosing among $h$ risky assets whose continuously compounded returns are given by the vector $\mathbf{r}_{t}^{s} \equiv\left(\begin{array}{lll}r_{1 t} & r_{2 t} \ldots & r_{h t}\end{array}\right)^{\prime}$. Portfolio weights are collected in the vector $\boldsymbol{\omega}_{t} \equiv\left(\omega_{1 t} \omega_{2 t} \ldots \omega_{h t}\right)^{\prime}$. The portfolio selection problem solved by a buy-and-hold investor with initial unit wealth becomes ${ }^{5}$

$$
\begin{gather*}
\max _{\boldsymbol{\omega}_{t}} E_{t}\left[U\left(W_{t+T}\left(\boldsymbol{\omega}_{t}\right) ; \boldsymbol{\theta}\right)\right] \\
\text { s.t. } W_{t+T}\left(\boldsymbol{\omega}_{t}\right)=\left\{\boldsymbol{\omega}_{t}^{\prime} \exp \left(\mathbf{R}_{t+T}^{s}\right)\right\} \tag{1}
\end{gather*}
$$

where $\mathbf{R}_{t+T}^{s} \equiv \mathbf{r}_{t+1}^{s}+\mathbf{r}_{t+2}^{s}+\ldots+\mathbf{r}_{t+T}^{s}$ is the vector of continuously compounded portfolio returns over the $T$-period investment horizon, and portfolio hares sum to 1 . Accordingly, $\exp \left(\mathbf{R}_{t+T}^{s}\right)$ is a vector of cumulated portfolio returns. No short-selling can be imposed through the constraint $\omega_{i t} \in[0,1]$ for $i=1,2, \ldots, h .{ }^{6}$

We approximate a Von-Neumann Morgenstern expected utility function with a function of four moments of the wealth distribution, of the form:

$$
\begin{equation*}
\hat{E}_{t}\left[U^{m}\left(W_{t+T} ; \boldsymbol{\theta}\right)\right]=\sum_{n=0}^{m} \kappa_{n} E_{t}\left[\left(W_{t+T}-v_{T}\right)^{n}\right] \tag{2}
\end{equation*}
$$

with $\kappa_{0}>0$, and $\kappa_{n}$ positive (negative) if $n$ is odd (even). When $n=2$, the investor has mean-variance preferences $(M V)$ : under non-satiation and risk aversion, marginal utility is positive $\left(U^{\prime}>0\right)$ and decreasing $\left(U^{\prime \prime}<0\right)$ in wealth. Assuming decreasing absolute risk aversion, we further have $U^{\prime \prime \prime}>0$ (investors prefer positive skew) while, as shown by Kimball (1993), decreasing absolute prudence implies that $U^{\prime \prime \prime \prime}<0$.

### 2.2.1. The Return Process

We assume that the vector of continuously compounded returns, $\mathbf{r}_{t}=\left(r_{1 t}, r_{2 t}, \ldots, r_{h t}\right)^{\prime}$, is generated by a Markov switching vector autoregressive process driven by a common unobservable state variable, $S_{t}$, that takes integer values between 1 and $k$ :

$$
\begin{equation*}
\mathbf{r}_{t}=\boldsymbol{\mu}_{s_{t}}+\sum_{j=1}^{p} \mathbf{A}_{j, s_{t}} \mathbf{r}_{t-j}+\varepsilon_{t} \tag{3}
\end{equation*}
$$

Here $\boldsymbol{\mu}_{s_{t}}=\left(\mu_{1 s t}, \ldots, \mu_{h s_{t}}\right)^{\prime}$ is a vector of intercepts in state $s_{t}, \mathbf{A}_{j, s_{t}}$ is an $h \times h$ matrix of autoregressive coefficients associated with the $j$ th lag in state $s_{t}$, and $\varepsilon_{t}=\left(\varepsilon_{1 t}, \ldots, \varepsilon_{h t}\right)^{\prime} \sim N\left(\mathbf{0}, \boldsymbol{\Omega}_{s_{t}}\right)$ is a vector of Gaussian

[^3]return innovations with zero mean vector and state-dependent covariance matrix $\boldsymbol{\Omega}_{s_{t}}$ :
$$
\boldsymbol{\Omega}_{s_{t}}=E\left[\left(\mathbf{r}_{t}-\boldsymbol{\mu}_{s_{t}}-\sum_{j=1}^{p} \mathbf{A}_{j, s_{t}} \mathbf{r}_{t-j}\right)\left(\mathbf{r}_{t}-\boldsymbol{\mu}_{s_{t}}-\sum_{j=1}^{p} \mathbf{A}_{j, s_{t}} \mathbf{r}_{t-j}\right)^{\prime} \mid s_{t}\right]
$$

The state-dependence of the covariance matrix captures the possibility of heteroskedastic shocks to asset returns, which is supported by strong empirical evidence, c.f. Bollerslev et al.(1992). Each state is assumed to be the realization of a first-order, homogeneous Markov chain as the transition probability matrix, $\mathbf{P}$, governing the evolution in the common state variable, $S_{t}$, has elements

$$
\begin{equation*}
\mathbf{P}[i, j]=\operatorname{Pr}\left(s_{t}=j \mid s_{t-1}=i\right)=p_{i j}, \quad i, j=1, . ., k \tag{4}
\end{equation*}
$$

Conditional on knowing the state next period, the return distribution is Gaussian. However, since future states are never known in advance, the return distribution is a mixture of normals with the mixture weights reflecting the current state probabilities and the transition probabilities.

Even in the absence of autoregressive terms, (3)-(4) imply time-varying investment opportunities. For example, the conditional mean of asset returns is an average of the vector of mean returns, $\boldsymbol{\mu}_{s_{t}}$, weighted by the filtered state probabilities $\left(\operatorname{Pr}\left(s_{t}=1 \mid \mathcal{F}_{t}\right), . ., \operatorname{Pr}\left(s_{t}=k \mid \mathcal{F}_{t}\right)\right)^{\prime}$, conditional on information available at time $t, \mathcal{F}_{t}$. Since these state probabilities vary over time, the expected return will also change. In addition, this setup can readily be extended to incorporate a range of predictor variables such as the dividend yield. This is done simply by expanding the vector $\mathbf{r}_{t}$ with additional predictor variables, $\mathbf{z}_{t}$, and modeling the joint process $\mathbf{y}_{t}=\left(\begin{array}{ll}\mathbf{r}_{t}^{\prime} & \mathbf{z}_{t}^{\prime}\end{array}\right)^{\prime}$.

When regimes are persistent and mean returns, $\boldsymbol{\mu}_{s_{t}}$, differ across states, expected returns vary over time. Similarly, when the covariance matrices, $\boldsymbol{\Sigma}_{s_{t}}$, differ across states, there will be predictability in higher order moments such as volatilities, correlations, skews and tail thickness. Predictability is therefore not confined to mean returns but carries over to the entire return distribution. Effectively, the return distribution is calculated as a weighted average of the individual, state-specific distributions using weights that are updated as new return data arrive.

### 2.2.2. The Portfolio Allocation Problem

We now indicate how to solve the investor's optimal asset allocation problem when preferences are defined over moments of terminal wealth (2) and returns follow the regime switching process (3)-(4). We follow Guidolin and Timmermann (2008) and compute the expected utility as

$$
\hat{E}_{t}\left[U^{m}\left(W_{t+T} ; \boldsymbol{\theta}\right)\right]=\sum_{n=0}^{m} \kappa_{n} \sum_{j=0}^{n}(-1)^{n-j} v_{T}^{n-j}\binom{n}{j} E_{t}\left[\left(\boldsymbol{\omega}_{t}^{\prime} \exp \left(\mathbf{R}_{t+T}^{s}\right)\right)^{j}\right] .
$$

In turn, the $n$th moment of the cumulated return on the portfolio is given by:

$$
E_{t}\left[\left(\boldsymbol{\omega}_{t}^{\prime} \exp \left(\mathbf{R}_{t+T}^{s}\right)\right)^{n}\right]=\sum_{n_{1}=0}^{n} \cdots \sum_{n_{h}=0}^{n} \lambda\left(n_{1}, n_{2}, \ldots, n_{h}\right)\left(\prod_{i=1}^{h} \omega_{i}^{n_{i}}\right) M_{t+T}^{(n)}\left(n_{1}, \ldots, n_{h}\right)
$$

where $\sum_{i=1}^{h} n_{i}=n, 0 \leq n_{i} \leq n(i=1, \ldots, h)$,

$$
\lambda\left(n_{1}, n_{2}, \ldots, n_{h}\right) \equiv \frac{n!}{n_{1}!n_{2}!\ldots n_{h}!}
$$

and $M_{t+T}^{(n)}\left(n_{1}, \ldots, n_{h}\right)$ can be evaluated recursively, using equations in the Appendix. The moments of the wealth distribution can thus be obtained by solving a small set of difference equations corresponding to the number of regimes in the return distribution. The otherwise complicated numerical problem of optimal asset allocation is reduced to one of solving for the roots of a low-order polynomial. This solution is closed-form in the sense that it is computable with a finite number of elementary operations.

In our application below we use $m=4$ moments in the preference specification. The weights on the first four moments of the wealth distribution are determined to ensure that our results can be compared to those in the existing literature that uses power utility functions. For a given coefficient of relative risk aversion, $\theta$, the power utility function serves as a guide in setting values of $\left\{\kappa_{n}\right\}_{n=0}^{m}$ in (2). Expanding the powers of $\left(W_{t+T}-v_{T}\right)$ and taking expectations, we obtain the following expression for the four-moment preference function:

$$
\begin{equation*}
\hat{E}_{t}\left[U^{4}\left(W_{t+T} ; \theta\right)\right]=\kappa_{0, T}(\theta)+\kappa_{1, T}(\theta) E_{t}\left[W_{t+T}\right]+\kappa_{2, T}(\theta) E_{t}\left[W_{t+T}^{2}\right]+\kappa_{3, T}(\theta) E_{t}\left[W_{t+T}^{3}\right]+\kappa_{4, T}(\theta) E_{t}\left[W_{t+T}^{4}\right] \tag{5}
\end{equation*}
$$

where $^{7}$

$$
\begin{aligned}
\kappa_{0, T}(\theta) & \equiv v_{T}^{1-\theta}\left[(1-\theta)^{-1}-1-\frac{1}{2} \theta-\frac{1}{6} \theta(\theta+1)-\frac{1}{24} \theta(\theta+1)(\theta+2)\right] \\
\kappa_{1, T}(\theta) & \equiv \frac{1}{6} v_{T}^{-\theta}[6+6 \theta+3 \theta(\theta+1)+\theta(\theta+1)(\theta+2)]>0 \\
\kappa_{2, T}(\theta) & \equiv-\frac{1}{4} \theta v_{T}^{-(1+\theta)}[2+2(\theta+1)+(\theta+1)(\theta+2)]<0 \\
\kappa_{3, T}(\theta) & \equiv \frac{1}{6} \theta(\theta+1)(\theta+3) v_{T}^{-(2+\theta)}>0 \\
\kappa_{4, T}(\theta) & \equiv-\frac{1}{24} \theta(\theta+1)(\theta+2) v_{T}^{-(3+\theta)}<0 .
\end{aligned}
$$

Notice that the expected utility from final wealth increases in $E_{t}\left[W_{t+T}\right]$ and $E_{t}\left[W_{t+T}^{3}\right]$, so that higher expected returns and more right-skewed distributions lead to higher expected utility. Conversely, expected utility is a decreasing function of the second and fourth moments of the terminal wealth distribution.

A solution to the optimal asset allocation problem can now easily be found from (5) by solving a system of cubic equations in $\hat{\boldsymbol{\omega}}_{t}$ derived from the first and second order conditions

$$
\left.\nabla_{\boldsymbol{\omega}_{t}} \hat{E}_{t}\left[U^{4}\left(W_{t+T} ; \theta\right)\right]\right|_{\hat{\boldsymbol{\omega}}_{t}}=\mathbf{0}^{\prime},\left.\quad H_{\boldsymbol{\omega}_{t}} \hat{E}_{t}\left[U^{4}\left(W_{t+T} ; \theta\right)\right]\right|_{\hat{\boldsymbol{\omega}}_{t}} \text { is negative definite. }
$$

Thus $\hat{\boldsymbol{\omega}}_{t}$ sets the gradient, $\nabla \boldsymbol{\omega}_{t} \hat{E}_{t}\left[U^{4}\left(W_{t+T} ; \theta\right)\right]$, to a vector of zeros and produces a negative definite Hessian matrix, $H_{\omega_{t}} \hat{E}_{t}\left[U^{4}\left(W_{t+T} ; \theta\right)\right]$.

[^4]Mean-Variance-Skew preferences are given by:

$$
\begin{equation*}
\hat{E}_{t}\left[U^{3}\left(W_{t+T} ; \theta\right)\right]=\kappa_{0, T}(\theta)+\kappa_{1, T}(\theta) E_{t}\left[W_{t+T}\right]+\kappa_{2, T}(\theta) E_{t}\left[W_{t+T}^{2}\right]+\kappa_{3, T}(\theta) E_{t}\left[W_{t+T}^{3}\right] \tag{6}
\end{equation*}
$$

where now $\kappa_{0, T}(\theta) \equiv v_{T}^{1-\theta}\left[(1-\theta)^{-1}-1-\frac{1}{2} \theta-\frac{1}{6} \theta(\theta+1)\right], \kappa_{1, T}(\theta) \equiv v_{T}^{-\theta}\left[1+\theta+\frac{1}{2} \theta(\theta+1)\right]>0, \kappa_{2, T}(\theta) \equiv$ $-\frac{1}{2} \theta v_{T}^{-(1+\theta)}(2+\theta)<0$, and $\kappa_{3, T}(\theta) \equiv \frac{1}{6} \theta(\theta+1) v_{T}^{-(2+\theta)}>0$.
while MV preferences simplify to:

$$
\begin{equation*}
\hat{E}_{t}\left[U^{2}\left(W_{t+T} ; \theta\right)\right]=\kappa_{0, T}(\theta)+\kappa_{1, T}(\theta) E_{t}\left[W_{t+T}\right]+\kappa_{2, T}(\theta) E_{t}\left[W_{t+T}^{2}\right] \tag{7}
\end{equation*}
$$

where $\kappa_{0, T}(\theta) \equiv v_{T}^{1-\theta}\left[(1-\theta)^{-1}-1-\frac{1}{2} \theta\right], \kappa_{1, T}(\theta) \equiv v_{T}^{-\theta}(1+\theta)>0$, and $\kappa_{2, T}(\theta) \equiv-\frac{1}{2} \theta v_{T}^{-(1+\theta)}<0$.
Thus the optimal portfolio composition, which in a standard MV problem with Gaussian return depends only on the variance-covariance matrix of returns and on risk aversion, also depends on:

1. differences between mean returns, $\mu_{1}, \mu_{2}$, and variances, $\sigma_{1}, \sigma_{2}$, (and more generally covariance parameters) across states. For example, skew in the return distribution can only be induced provided that $\mu_{1} \neq \mu_{2}$, c.f. Timmermann (2000).
2. The current state probabilities $\left(\pi_{t}, 1-\pi_{t}\right)$ which determine moments of returns at all future points provided that either the mean or variance parameters differ across states ( $\mu_{1} \neq \mu_{2}$ or $\sigma_{1} \neq \sigma_{2}$ ).
3. State transition probabilities which also affect the speed of mean reversion in the investment opportunity set towards its steady state.
4. The number of moments of the wealth distribution that matters for preferences, $m$, in addition to the weights on the various moments.
5. The investment horizon, $T$.

Our benchmark results assume that $\theta=5$, a coefficient of relative risk aversion compatible with the bulk of empirical evidence. Later we present robustness results that allow this coefficient to assume both larger and smaller values.

## 3. Empirical Results

### 3.1. Data and Descriptive Statistics

We analyze three datasets of monthly equity returns. The first one - which we label IE - comprises eight international portfolios, three of which are from emerging markets (1988:01-2008:08). The second one covers ten US industry portfolios (IND), from July 1926 to July 2008, while the third one refers to ten Book-toMarket (BM) sorted portfolios from five geographical areas plus the world market portfolios (1975:012007:12). Our choice of data, if anything, distorts against finding a large extent of non-normalities. Indeed, our focus is neither on individual security returns nor on data sampled at higher frequencies, where it is easier to uncover non-normal features. We also avoid analyzing portfolios of size and momentum portfolios, let
alone hedge-fund returns, that are already known to display asymmetries that are exploitable in a portfolio setting (see e.g. Guidolin and Nicodano, 2008; Hong et al., 2007).

Table 1 reports summary statistics, with the lower parts displaying the correlation/co-kurtosis matrix and the co-skewness matrix respectively.

Panel A shows a wide dispersion of monthly mean returns and volatilities across international equity portfolios. Also the ratio of expected return to volatility (Sharpe) covers a wide range, from -0.064 and 0.059 for Japan and EM Asia, to 0.193 and 0.174 respectively for EM Europe \& Middle East and EM Latin America. Correlations involving the UK and North America are generally higher than those involving Emerging Markets, whose cross-correlations never exceed 0.491. ${ }^{8}$

On the contrary, monthly mean returns are not particularly disperse in the US Industry dataset, ranging from 0.831 for Telecommunication, to 1.097 for Energy (Panel B). Yet their Sharpe ratios differ markedly, from 0.095 for Other Industries to 0.143 for Non Durables. Industry portfolios unsurprisingly display higher cross-correlations, always in excess of 0.5 , than other international portfolios being exposed to the same country risk factor.

Dispersion in mean returns and Sharpe ratios is also high in Book-to-Market international portfolios (Panel C). As in Fama and French (1998), higher mean returns on value portfolios are the norm, with the exception of the US market, with a peak of 1.73 for the UK one. Value portfolios usually yield the highest Sharpe ratios as well, e.g. 0.201 for the United Kingdom Value. There are large and significant correlations between Value and Growth portfolios within country. For instance, the correlation between EU ex-UK ex-Scand Value and EU ex-UK ex-Scand Growth is 0.850 .

An investor with higher-order preferences cares about higher order moments and co-movements in returns. This is why we also measure the co-skew of a triplet of stock returns $i, j, l=1, \ldots, h$ as in Jondeau and Rockinger (2004):

$$
\begin{equation*}
S_{i, j, l} \equiv \frac{E\left[\left(r_{i t}-E\left[r_{i t}\right]\right)\left(r_{j t}-E\left[r_{j t}\right]\right)\left(r_{l t}-E\left[r_{l t}\right]\right)\right]}{\left\{E\left[\left(r_{i t}-E\left[r_{i t}\right]\right)^{2}\right] E\left[\left(r_{j t}-E\left[r_{j t}\right]\right)^{2}\right] E\left[\left(r_{l t}-E\left[r_{l t}\right]\right)^{2}\right]\right\}^{1 / 2}} . \tag{8}
\end{equation*}
$$

When $i=j=l, S_{i, j, l}$ reduces to the third central moment of returns on asset $i$, which captures the traditional measure of skew, $S k e w_{i}=S_{i, i, i} / \sigma_{i}^{3}$. reported in the upper panel of Table 1. When $i \neq j \neq l, S_{i, j, l}$ gives a signed measure of the strength of the linear association among deviations of returns from their means across triplets of asset returns. A risk-averse investor dislikes negative values of $S_{i, j, l}$ corresponding to cases when returns in different markets are below their mean at the same time.

When only the returns on two assets are involved, $S_{i, j, j}$ reflects the strength of the linear association between squared deviations from the mean and signed deviations from the mean for a pair of assets. A security $i$ with negative $S_{-i, i, i}$ coefficients for the majority of all possible pairs of returns on other securities (denoted as $-i$ ) is a security that becomes highly volatile when other securities give low returns, and viceversa. To a risk averse investor this is an unattractive feature since risk rises in periods with low returns. A security $i$ with predominantly negative $S_{i,-i,-i}$ coefficients pays low (high) returns when other securities become highly volatile; again this feature is harmful to a risk-averse investor since the security performs

[^5]poorly when other assets are highly risky. The bottom panel of Table 1 reports the elements $S_{-i, i, i}$ and $S_{i,-i,-i}$ respectively above and below the diagonal.

Turning to fat tails in the return distribution, the co-kurtosis of a set of four stock returns $i, j, l, q=1, \ldots, h$ is equal to:

$$
\begin{equation*}
K_{i, j, l, b} \equiv \frac{E\left[\left(r_{i t}-E\left[r_{i t}\right]\right)\left(r_{j t}-E\left[r_{j t}\right]\right)\left(r_{l t}-E\left[r_{l t}\right]\right)\left(r_{q t}-E\left[r_{q t}\right]\right)\right]}{\left\{E\left[\left(r_{i t}-E\left[r_{i t}\right]\right)^{2}\right] E\left[\left(r_{j t}-E\left[r_{j t}\right]\right)^{2}\right] E\left[\left(r_{l t}-E\left[r_{l t}\right]\right)^{2}\right] E\left[\left(r_{q t}-E\left[r_{q t}\right]\right)^{2}\right]\right\}^{1 / 2}} . \tag{9}
\end{equation*}
$$

When $i=j=l=q, K_{i, j, l, q}$ becomes proportional to the coefficient of kurtosis, Kurt $_{i}=K_{i, i, i, i} / \sigma_{i}^{4}$ reported in the upper part of Table 1. When $i \neq j \neq l \neq q, K_{i, j, l, q}$ gives a signed measure of the strength of the linear association among deviations of returns from their means across four-tuples of asset returns. The term $K_{i, i, j, j}$, which is present in the middle section of Table 1, sheds light on the correlation between volatility shocks across markets. Large positive values are undesirable, reflecting that volatility tends to be large at the same time in market $i$ as in other market, thus increasing the overall portfolio risk. ${ }^{9}$

We can now comment on the relevance of higher order moments in our sample. The Jarque-Bera test referring to the IE dataset (Panel 1) rejects Normality at $1 \%$ for five equity portfolios, but it cannot reject the null for two of them. ${ }^{10}$ Only two (four) values of the co-kurtosis (co-skewness) differ statistically from zero at $1 \%$ significance level. On the contrary, normality is rejected at the $1 \%$ level for all equity portfolios in the IND and BM datasets, described in Panel B and C respectively. Co-kurtosis is statistically different from zero at $1 \%$ significance level in all (most) portfolios, while co-skew is (almost) always negligible in the IND (BM) datasets. Thus it appears that non-normalities are moderate in the first dataset, intermediate in Book-to-Market sorted portfolios and substantial in the industry portfolios. However, systematic skewness appears in the BM dataset only.

### 3.2. The Return Generating Process

Empirical analyses of portfolio problems often specify an exogenous distribution of returns. We instead perform several specification tests, allowing our data to endogenously determine the number of regimes $k$. Table 2 reports the results of these tests, for up to $k=4$ regimes. We also let the test determine whether there should be an autoregressive term of order $p=1$, as opposed to no lag $(p=0)$. In Table 2, the term $\operatorname{MSIA}(1,1)$ is the same as $\operatorname{VAR}(1)$ - indicating linear predictability of returns - while MSIA( 1,0 ) indicates the Gaussian model with unpredictable returns. $\operatorname{MSIA}(2,1)$ allows for an autoregressive term of order 1 as well as for two regimes, while $\operatorname{MSIAH}(2,1)$ adds heteroskedasticity in the error terms in the form of a regime-switching covariance matrix for returns.

The information criteria do not discriminate between alternative return processes in the case of the first dataset (Panel A), confirming that non-normalities are modest. For the industry dataset, the BIC and HQ

[^6]criteria indicate that a $\operatorname{MSIH}(2,0)$ return process provides a good fit of the data. Other well-fitting models include $\operatorname{MSIH}(3,0)$ and $\operatorname{MSIH}(4,0)$, but we stick to the more parsimonious specification. ${ }^{11}$

We can now turn to Table 3, which displays estimates of the parameters of the Gaussian i.i.d. return process in Panel A and of the two-regime model in Panel B. In our description below, we focus on the two-regime representation, as the parameters associated with the single-state one mirror closely the sample descriptive statistics. In one state - which we name Bear - most equity portfolios have lower mean returns. The bear regime tends to last from a minimum of 3.43 months in the BM dataset to a maximum of 5.60 month in the Industry dataset (see Panel C of tables 3A, 3B and 3C). The persistence of bull markets is always higher than the bear one- consistent with business cycle evidence - and it is highest in the Industry dataset (19.46 months) and lowest in the International one ( 9.7 months). The conditional correlation/volatilities matrixes are estimated with high precision under the two-state model, with a level of significance almost always exceeding $1 \%$. Volatilities in the bear state are always larger than in the bull state, with the exception of EM Latin America in International database ${ }^{12}$ Finally, mean returns usually (do not) differ from zero in the (bear) bull state. These regularities sharply differentiate the two regimes in all datasets.

In Table 3A, the two-regime representation leads to even greater dispersion across International and Emerging equity portfolios than that already present in the single state representation. In the Bear regime, four out of eight markets have negative Sharpe ratios (JP, Pacific EX JP, Europe EX UK, UK), whereas Emerging Markets still have positive Sharpe ratios with EM Latin America displaying a particularly high Sharpe ratio of 0.428 . North American stocks turn out to have the second highest Sharpe ratio, 0.090 , with a relatively low volatility of 4.28 . Under the Bull regime, volatilities are lower for every stock market but for Emerging Latin America, which is more volatile during Bull than Bear states. Nonetheless, the Sharpe ratio of EM Latin America (and also EM Europe \& Middle East) exceeds 0.38, far higher than others - all of which are positive. Correlations involving North American and EU-ex-UK portfolios tend to be higher in bear markets, confirming the insight by Longin and Solnik (2001) that diversification appears more difficult in bear states. Panel C shows that the Bull regime is almost twice as likely as the Bear one ( 0.348 vs. 0.652 ).

For industry portfolios in Table 3B, the Bull regime is more than three times as likely as the Bear one (0.224 vs. 0.776 ). ${ }^{13}$ In line with positive mean returns of North American stocks in the IE dataset, every mean US Industry return is positive in the Bear regime, with a wide variation between Energy and Durables on the one side, with respectively 0.760 and 0.751 , and Other and Shops at the other end of the spectrum, with 0.111 and 0.081 respectively. Looking at volatilities, we note that Energy has a relatively low volatility (9.318) compared with that of e.g. Durables (13.181). Note that both correlations and volatilities are higher for industry than for country portfolios in the bear regime, suggesting that country diversification is more powerful than industry diversification, as found by Griffin and Karolyi (1998).

Under the Bull regime, the ranking of industry portfolios changes. Considering Sharpe ratios, the highest

[^7]are Non Durables,Telecommunications and Health, all exceeding 0.284. All volatilities are lower under Bull than under Bear regimes, while correlations are comparable across stock portfolio, with correlations being higher in Bull rather than Bear regimes - even for defensive industries like Utilities. ${ }^{14}$

In Book-to-Market international portfolios, EU ex-UK ex-Scand Value has the highest Sharpe ratio (0.448). Bull-market correlations across Value portfolios range from 0.319 (United States/Scandinavia) to 0.562 (EU ex-UK ex-Scand/UK). Correlations across Growth portfolios are generally higher, ranging from 0.327 (United States/Asia \& Pacific) to 0.626 (EU ex-UK ex-Scand/United States). Conditional correlations are similar to those of the single state model.

When Bear, UK Value and Asia \& Pacific Value turn out to have the highest Sharpe ratios, respectively 0.246 and 0.222 ; whereas EU ex-UK ex-Scand and Asia \& Pacific growth the lowest, -0.035 and -0.018 .

Ang and Chen (2002) study the changing correlation across bull and bear states in US BM portfolios. We do confirm their result that US Value has higher correlation with the world portfolio than US Growth in bear markets. However, the same finding does not carry over to other international BM portfolios: value and growth stocks display a similar pattern of correlations across states, with all correlations being larger in bear states. Petkova and Zhang (2005) argue that the beta risk of value-minus-growth is higher in bear states: given that Value stocks have far higher volatilities than Growth stocks in bear markets, their finding carries over to our international dataset. This also rationalizes why the expected return differential between value and growth stocks appears to be largest in Bear markets. For instance, the difference in mean returns between Asia \& Pacific Value (1.168) and Asia \& Pacific Growth (0.612) is 0.556 when Bull, increasing to 1.741 in Bear markets when mean returns respectively equal 1.623 and $-0.118 .{ }^{15}$

Against this background, we can now study how these equity indexes enter optimal portfolios.

### 3.3. The Composition of Optimal Equity Portfolios

The left-hand side panel of Table 5 reports portfolio composition when short sales are not possible, which we comment upon in this section. The panel is further divided into columns according to the investor horizon. Each page is divided horizontally according to investors' preferences. In the upper panel there is the classic $M V(1)$ case, followed by the unconditional allocations when there are two regimes $k=2$ and preferences are defined over $m=2,3$ and 4 moments. The lowest two parts of each table describe portfolio allocations conditional on the two states. In this commentary, we will mostly comment on the classic mean variance case, on the case where the mean variance investor predicts changing moments (MV(2)) and on four moment preferences, $\operatorname{MVSK}(2)$.

### 3.3.1. International

A classic mean-variance investor cares only about the trade-off between means and volatilities and invests only in those portfolios with the highest Sharpe ratios, for given correlation. As to the International data set,

[^8]the MV(1) investor weights only EM Latin America (from 0.383 to 0.415 depending on $T$ ) and EM Europe \& Middle East (from 0.617 to 0.585 ) which provide the highest Sharpe ratios across every investment horizon and exhibit a relatively low cross-correlation (0.479).
When allowing for regimes, a mean-variance investor weighs positively North America stocks (from 0.413 to 0.243 ) along with Latin America and Middle East. This occurs unconditionally, but it is especially true in the Bear state, where portfolio composition is shifted away from emerging markets towards North America (from 0.997 to 0.690 , depending on the investor horizon $T$ ). In line with findings of Guidolin and Timmermann (2008), this is explained by the high Sharpe ratio of this country index in Bear states. Portfolio composition does not significantly differ between Bear and Bull states in other respects, with only North America receiving more weight when Bear and Latin America and Middle East when Bull.

We observe a further increase in diversification when the investor has four-moment preferences, MVSK. In fact, weights are positive for every country index when the investment horizon is one month and for 7 out of 8 when the horizon reaches one year. The co-kurtosis between stock portfolios helps rationalizing portfolio weights: Japan, which receives a weight of 0.253 in the Bull state, has the lowest co-kurtosis with EM EU \& Middle East (1.253), which in turn has the second highest weight (0.239). On the contrary, Pacific ex-Japan and EM Asia, which display the highest co-kurtosis (2.833), are the portfolios with the lowest weights, respectively 0 and 0.035 .

Not only do MVSK preferences lead to the highest level of diversification, but also to the most stable portfolio holding dynamics. Indeed, comparing MV(2) with MVSK holdings in Figure 2, we see that the two-regime mean-variance model is characterized by infrequent spikes, which, on the contrary, do not appear in the case of MVSK preferences. Additionally, portfolio weights for classic mean-variance investors MV(1) are more stable than $\operatorname{MV}(2)$ ones, as $\mathrm{MV}(1)$ investors update expected returns and volatility estimates only slightly when including new data.

When the horizon increases up to $\mathrm{T}=120$, the composition of $\mathrm{MV}(1)$ investors does not vary significantly. On the contrary, MV(2) and MVSK investors concentrate their portfolios in three countries only, namely North America, EM Latin America and EM Europe and Middle East in roughly the same proportions. This indirectly suggests that non-normalities, which were moderate according to sample based tests, disappear as the horizon lengthens. ${ }^{16}$

### 3.3.2. Industry

Turning to the Industry data set, the optimal portfolio of a classic mean-variance investor, MV(1), is similar to that of an investor who also accounts for regimes, MV(2). In fact, both of them invest in Energy, Health and Non-Durables sectors, which have the highest Sharpe ratios. Furthermore the correlation between Energy and Health does not exceed 0.5. Considering separately the Bear and Bull state, we note that Energy is overweighted in Bear states, as it has the highest mean return and a low volatility, whereas positions in Health and Non-Durables increase in Bull markets when they have the highest Sharpe ratios.

When allowing for preferences on third and fourth moments (MVSK), the level of the portfolio diversification slightly increases. Investors hold Energy, Health, Telecommunications and Durables, probably due to

[^9]their low co-kurtosis (e.g. 2.744 between Energy and Telecommunications) and the positive co-skew between Energy and Durables (0.096). Differently from MV investors, MVSK ones choose almost the same allocation between the two regimes, with only Energy and Telecommunications receiving slightly higher weights when Bear, as opposed to Health and Non-Durables when Bull.

As to the portfolio holding dynamics, MV(1) has the most stable one, whereas both MV(2) and MVSK have infrequent spikes. The inclusion of higher moments reduces the size of the spikes, which is lower for MVSK.

Increasing the investment horizon to $\mathrm{T}=120$ alters the portfolio composition of MV(1) investors, tilting it towards Telecommunications and away from Health. Also MV(2) investors tilt their portfolio allocation away from Health, but towards Non Durables. The MVSK allocation lies in the middle of these two patterns, as it overweights Non-Durables and underweights both Health and Telecommunications.

### 3.3.3. Book-to-Market

Turning now to our last dataset, portfolio weights for a MV investor are concentrated in Value stocks for $T=1,12$ - namely UK Value and Scandinavia Value - irrespective of regime considerations.

In the Bull regime, both EU Ex-UK ex-Scandinavia Value in $T=1$ and United States Growth for $T>1$ receive a large weight having a very large conditional Sharpe ratio ( 0.448 and 0.497 respectively). When Bear, the first portfolio receives a zero weight having a conditional Sharpe ratio of 0.006 , while UK Value ( 0.246 ) is overweighted and Scandinavia (0.192) and US Growth ( 0.188 ) are also included in the investor's portfolios.

The highest degree of diversification is achieved when the investor has four-moment preferences (MVSK). Specifically, an investor who accounts also for skewness and kurtosis would diversify in 6 on 11 portfolios, with Value stocks being more heavily weighted. This behavior is not only explained by the low correlations between Asia \& Pacific Value and UK Value (0.399) and between Asia \& Pacific Value and Scandinavia Value ( 0.370 ) but also by the low co-kurtosis - for instance 1.833 between UK Value and Scandinavia Value and 1.902 between Asia \& Pacific Value and Scandinavia Value.

When Bear, 7 portfolios are included, with Growth stocks receiving higher weight with respect to the ergodic case, since they have much lower volatility than Value stocks. Such high degree of diversification holds in the Bull state as well.

Turning to the dynamics of portfolio weights, there is a striking difference between MV(1) and MV(2) investors in the stability of their holdings. Indeed, MV(1) appear to have more stable weights, due to the absence of spikes that, on the contrary, characterize MV $(2)$ holdings - especially in the long run. Furthermore, portfolio holdings are more stable with MVSK than with the other models, despite the presence of some spikes of small size.

When the horizon increase to $T=60$, US Value and US Growth increase in importance to totally replace value portfolios for $T=120-$ when returns are considered gaussian. This pattern is present but less marked with MVSK model: when $T=120$ the portfolio composition involves also Scandinavia Value, along with US Value and US Growth.

### 3.3.4. Co-Skew and Co-Kurtosis Properties

Table 4 reports the $T=1$ steady-state skew and kurtosis properties of each stock portfolio, implied by the Markov-switching return process. We can first observe that the higher order moments implied by our regime-switching model closely match their sample counterparts, offering further evidence that the model is not misspecified. We can also investigate whether higher order properties explain why some stocks enter prominently in MV portfolios but have marginal, if any, role in MVS and MVSK portfolios. Candidates are both EM Latin America and EM EU \& Middle East in the IE dataset.

Table 4 shows that the EM Latin America portfolio has large negative values of both own-market skew ( $S_{U S, U S, U S}$ ), and co-skews $S_{U S, U S, j}, S_{U S, j, j}$, producing either the largest negative or second largest negative sample estimates of these moments across all regions. Hence EM Latin America stock returns tend to be negative when volatility is high in other markets and they are more volatile when other markets experience negative returns. The implication is that they provide little or no hedge against adverse return or volatility shocks in other markets. A similar limitation affects the desirability of EM EU \& Middle East stocks. These effects allow us to explain why aversion to skew in the distribution of final wealth reduces the weights of these stock portfolios.

In the Industry data set, the portfolio becomes concentrated in Energy when moving from MV to MVS preferences for $\mathrm{T}=1$. This can be traced back to the positive skewness of Energy returns ( 0.034 ) as well as to the negative skew of both Non-Durables $(-0.147)$ and Health ( -0.037 ). Co-skewness between portfolios are -0.092 and -0.067 (Energy/ Non-Durables), -0.010 and -0.012 (Energy/ Health) and -0.101 and -0.066 (Health/ Non-Durables). Thus, there is a similarity between EM Latin America and EM EU \& Middle East and Health/ Non-Durables, as well as for North American shares and Energy.

When considering MVSK, the importance of both own-kurtosis and co-kurtosis help increase the number of portfolio held. Indeed, Telecommunications and Energy, which are the most demanded, have the lowest own-kurtosis -respectively 5.761 and 5.315 - and the lowest co-kurtosis, 2.744.

In the BM dataset, we saw that investors with MVS preferences do not invest in the Scandinavian Value stocks that enter MV(1) portfolios together with UK Value. This can be probably traced back to the own implied skewness of this equity portfolio, which exceeds by far the one of UK Value ( 0.065 versus 0.0385 ), as well as to their negative co-skewness coefficients ( -0.068 and -0.049 ).

Table 4 shows that implied estimates of the co-kurtosis are highest for UK Value/UK Growth stock indexes (6.853). This region also produces high estimates of own-market kurtosis (6.919 and 10.34). The high value of own-market kurtosis for UK Value may explain why the allocation to this region does not increase further when shifting from preferences defined over skew $(m=3)$ to those over kurtosis as well ( $m=4$ ).

In general, stocks entering optimal portfolios of four-moment preferences investors have good co-kurtosis properties, with co-kurtosis coefficients ranging from 1.864 to 2.357 . For instance, Asia Pacific Value stocks - that are not demanded at all under MV preferences - enter optimal portfolios of both MVK and MVSK investors having a co-kurtosis of 1.902 with Scandinavian value and 2.357 with UK value. By comparison, US Growth, which dominates Asia Pacific Value in terms of implied Sharpe ratio ( 0.267 versus 0.184 ), has higher co-kurtosis ( 1.936 with Scandinavian and of 3.677 with UK Value).

### 3.3.5. Regularities across datasets

This subsection highlights the regularities found in all datasets.
A first regularity is that an investor with preferences defined over four moments (MVSK) always reaches the highest degree of diversification. A MV(1) investor has a concentrated portfolio, holding only two indexes in IE and Book-to-Market dataset and three indexes in the Industry one. A MV(2) investor respectively includes six, three and two portfolios in the IE, Industry and Book-to-Market datasets. On the contrary, under MVKS, five Industry, six Book-to-Market and every country index have positive weight. Portfolio holdings in MV models are focussed on portfolios having higher Sharpe ratios and lower correlations. Adding higher moment preferences as well results in the addition of those stock portfolios displaying lower implied co-kurtosis and higher co-skewness. However, increased diversification is especially attached to kurtosis aversion. This is clear from the comparison of MVS and MVK portfolios in the case of the International dataset: a strictly greater number of indexes receives non-zero portfolio weights when the investor has MVK instead of MVS preferences for every time horizon (except for $\mathrm{T}=120$ ). Thus, it appears that skewness aversion induces concentration in a subset of assets with good skewness properties - as already known from the literature. In the International dataset, MVS investors choose only the 3 highest Sharpe-ratio portfolios, giving more weight to EM EU and Middle East, which has the lowest skew. Moreover, co-skew between those three portfolios is moderate.

Another regularity concerns the volatility of portfolio holdings over time. Indeed, MV(2) models always entail more volatile weights, with infrequent spikes, whose size increases with increasing horizon. On the other hand, with MVSK, spikes vanish in IE dataset and are small in both Industry and Book-to-Market data. Therefore, transactions costs are more likely to adversely affect a mean-variance investor rather than a four-moment one and could possibly reverse the ranking of models in term of performance.

Ang and Bekaert (2004) suggest that RS strategies are relatively robust to transaction costs because they are designed to exploit changes in expected returns and volatilities that are associated with infrequent changes of regimes is relatively high. Our findings qualify this observation: the very contribution to the stability of portfolio shares is offered by the higher moments, given that they exhibit higher volatility under MV(2) than MV(1), i.e. when we allow for two regimes.

It is also the case that a shorter horizon increases the sensitivity of portfolio composition with respect to the current state of the market. For instance, a $\mathrm{T}=1$ investor in the Industry dataset weights heavily stock indexes that perform well in the bear state when the probability of being in a bear state is high, because the bear state is persistent. A MV investor thus gives 0.612 weight to the Energy portfolio for $\mathrm{T}=1$. On the contrary, a $\mathrm{T}=120$ investor, believing to be in a bear state, cares also about stock portfolios that outperform in bull markets as she knows that the chances of shifting to a bull regimes are higher. Non-Durable stocks are then weighted 0.516 , thanks to their attractiveness in bull states and despite their low mean return (0.162) in bear markets, while the weight on Energy falls to 0.253.

Last but not least, in each dataset there are several stock portfolios that are defensive in a traditional way, i.e. display relatively low correlations with other portfolios. Three portfolios that appear to be defensive for investors that recognize changes in mean returns and volatilities across regimes are North American, Energy and, to much lesser extent, US and Scandinavia Growth.

## 4. The Ex-Post Performance of Equity Portfolios

We know that the expected utility of an investor who cares about higher moments falls when $\mathrm{s} / \mathrm{he}$ overlooks predictability in returns and/or higher order moments of the return distribution. This is, for instance, the case in Ang and Bekaert (2002) in an international portfolio problem - provided that the asset menu includes a short-term bond allowing investors to abandon equities in the bear state. It is also the case when dealing with size-sorted equity portfolios, as in Guidolin and Nicodano (2008), due to the dismal performance of small caps in bear states. Out-of-sample analyses confirm that gains are large when diversifying internationally if the possibility of shifting into cash in bear states is allowed (Ang and Bekaert (2004) and Guidolin and Timmermann (2008)). In these papers the benchmark is rather extreme, being the mean-variance allocation with no predictability. However, we know that ex-post gains from timing both volatility (Fleming et al., 2001) and higher order moments (Jondeau and Rockinger, 2009) can be large. These papers find positive gains in a model where expected returns are constant and there are no regimes. We now turn to an assessment of out-of-sample performance gains in our three datasets, extending previous evidence along several dimensions.

We recursively estimate all the parameters of the models described in Table 0 and proceed to calculate the portfolio performance figures at all points in time. For the Markov-switching model this implies reestimating all parameters and the state probability vector on an expanding window of data using the EM algorithm. For other models, only the parameters are estimated recursively by MLE. The out-of-sample period for our International, Industry and Book-to-Market International data run from 1998:01-2008:07, 1980:01-2008:07, and 1995:01-2007:12 respectively. ${ }^{17}$

First, we use three different indicators of portfolio performance (see Table 6). The Sharpe ratio does not capture any effect on the skewness or kurtosis of wealth, while the Sortino ratio falls when downside risk increases. This enables to compare the performance of investors endowed with different preferences, by checking - for instance - whether the Sharpe/Sortino ratios of a 3 or 4 moment-preference investor exceeds the one of a MV investor. We can similarly analyze whether an investor with mean-variance preferences achieves better performance by considering time varying mean and variances across market regimes, i.e. by becoming what we label a $\mathrm{MV}(2)$ investor. And we also compute the certainty equivalent of maximum utility, CEQ, associated with different investor preferences ${ }^{18}$. Finally, Table 7 will report two other indicators of performance. The Treynor ratio indicates excess return on one unit of systematic risk, $\beta$, while Jensen's alpha captures excess-returns that are not associated to such systematic risk. It may well be that such measures, as well as the Sharpe ratio, increase with more negative skew in portfolio returns ${ }^{19}$, which would instead reduce welfare of a MVS investor.

A first observation concerns the relative performance of the equally weighted strategy vis-à-vis optimizing ones. Recently, the literature suggested that the equally weighted strategy would be the appropriate benchmark to evaluate the relative performance of active strategies (DeMiguel et al., 2009). From this point

[^10]of view, we observe that $1 / \mathrm{N}$ never consistently outperforms optimizing strategies. According to $5 \%-95 \%$ confidence bounds, the equally weighted strategy is always equivalent or dominated by the other ones. It achieves its best performance in the International dataset for a long time horizon ( $\mathrm{T}=120$ ), since it ranks first according to both Sharpe and Sortino Ratio. Yet, the other performance measures do not support this result even in this case ${ }^{20}$. In the other two datasets the equally weighted strategy is never the first best and, for $\mathrm{T} \geq 60$ months, it is almost always dominated according to all the five measures (Sharpe, Sortino and Treynor ratios, CEQ and Jensen's Alpha).

We now turn to the relative performance among optimizing strategies. Panel A of Tables 6 and 7 reveals that MV models are often the best performers in International diversification problems according to any measure. Such good ex-post performance of Mean Variance portfolios is perhaps not surprising, as we saw that non-normalities are moderate in this data-set. Furthermore, we already know from Ang and Bekaert (2004) that the out-of-sample cost of adopting i.i.d. mean-variance strategies is low when there is no risk-free rate. However, when $\mathrm{T}=12,60$, predicting regimes seems to matter, since MVS and MV(2) are either equivalent or dominate the simple $\mathrm{MV}(1)$ strategy. ${ }^{21}$ For instance, when $\mathrm{T}=60$, the 2-regimes strategies dominate MV(1) according to Sharpe Ratio, Treynor Ratio and Jensen's Alpha. In particular, unreported figures for $\mathrm{MV}(1)$ show that Jensen's $\alpha$ confidence interval is negative: $[-82.828,-58.636]$ while it is strictly positive for the other strategies. When $\mathrm{T}=12$, following a $\mathrm{MV}(2)$ strategy as opposed to a MV(1) delivers a CEQ of 4.993 rather than 4.713 and a strictly preferable Jensen's $\alpha(22.530$, with confidence bounds [7.441,37.445] vs. -22.495 with bounds [ $-38.883,-4.464]$ ). Thus, accounting for regimes is rewarding for longer horizons even in this dataset with moderate non-normalities. This complements prior results by Fleming et. al. (2001), who study a short-horizon mean-variance investor with daily data. It also adds to the evidence in Jondeau and Rockinger (2009), who use a DCC specification for the return process.

In the Industry dataset both MV and MVSK models rank high among optimizing strategies. In particular, MVSK - as well as $\mathrm{MV}(2)$ - perform better for shorter horizons ( $\mathrm{T} \leq 12$ ). For instance, MVSK yields a higher CEQ (13.755 vs. 10.867) and higher Jensen's alpha (124.699 vs. 73.429 ) than the MV(1) strategy for $\mathrm{T}=1$. Moreover, it ranks at least second for both $\mathrm{T}=1$ and $\mathrm{T}=12$ according to all the performance measures, even if no optimizing strategy is clearly the best on all counts ${ }^{22}$. The importance of predicting both stock market regimes and higher order moments fades away with longer time horizons. MV(1) ranks first according to most metrics for $\mathrm{T}=60$ and according to all of them for $\mathrm{T}=120$. This result is very strong for $\mathrm{T}=120$, since $\mathrm{MV}(1)$ dominates all the other strategies when we consider Treynor Ratios or Jensen's alphas (see Table 7 Panel B). Its 5\%-95\% confidence bounds, [1.373,1.839] (Treynor Ratio) and [49.970, 69.109] (Jensen's Alpha), are strictly greater than the ones for the MVS strategy, which ranks second, [1.033,1.197] and $[24.127,35.387]$. All in all, a MV strategy deals well with industry diversification irrespective of the investment horizon - despite substantial non-normalities both in our sample statistics as well as in the characterization of the return generating process.

[^11]By contrast, accounting for higher order preferences and regimes adds economic value when investing in Book-to-Market international portfolios, as evident from panel C in both Table 6 and 7. Both the MVS and the MVSK strategy outperform the MV(1) model. Even restricing attention to MV preferences, it turns out that the CEQ of a MV(2) investor always exceeds that of a MV(1) investor.

The welfare of an investor following the MV(2) strategy is lower than both MVSK (18.643 vs. 17.701) when $\mathrm{T}=1$ and MVS for $\mathrm{T}=12,60,120$ ( 8.586 vs. $7.328,12.868$ vs. 3.378 and 30.206 vs. 14.964 respectively). Note that for $\mathrm{T}=60$, 120, confidence bounds are not overlapping ( $[11.278,14.762]$ vs. [2.360, 4.575] and [29.524,30.906] vs. [14.135,15.881] respectively) and MVS strictly dominates MV(2) in a statistical sense. Hence, MV investors would prefer to delegate MVS(K) ones when dealing with Book-to-Market portfolios, opposite to what happens with the Industry ones. This is confirmed by the certainty equivalent of the best MV strategy, which is also lower than the $\operatorname{MVS}(\mathrm{K})$ one for all time horizons.

A final observation concerns the relevance of skewness for longer investment horizons. Indeed, the MVS strategy dominates MVSK for $\mathrm{T} \geq 60 .{ }^{23}$ For $T=60$, for instance, the Sharpe (Sortino) Ratio of a MVS investor, 0.571 (12.868), strictly exceeds the one of MVSK managers, 0.186 (3.603), who represent the second best alternative among the optimizing strategies. ${ }^{24}$.More generally, it appears that predicting skewness rather than kurtosis is more important in all datasets, since MVS performances are never dominated by MVSK ones.

### 4.1. Robustness

So far we commented on results referring to the case of a risk aversion coefficient equal to 5 and no short sale. We now turn to a situation when short sales are allowed. In this case, all three performance measures are generally lower than with short sales constraint, particularly for long horizons. This is not surprising, as the typical extreme long or short positions involved by short selling are able to exacerbate any misspecification or imprecise estimation, especially in the longer run. This result confirms previous research, which points out the importance of restricting the volatility of portfolios weights to achieve higher out-of-sample performance. (among others, DeMiguel et al., 2009, Diris et al. 2008, Jagannathan and Ma, 2003).

Turning to the ordering of portfolio choice models, we see that allowing for short sales does not alter the relative ranking of performance when $\mathrm{T}=1$ in International data according to the performance measures reported in Table 6. Specifically, MV(1) is still the best model according to both Sharpe ratio and Sortino ratio, whereas MVSK yields the highest certainty equivalent. Moreover, the relative performance of the second best and third best model is also consistent with the previous case. Results are mixed when T increases to 12 , with the ranking being the same only according to the Sharpe ratio, as MV(1) gives the highest risk-return trade-off also with short sales.

Industry data do not show the same patterns when we remove short sale constraints. When $\mathrm{T}=1$, the ranking changes both according to the Sharpe ratio and Certainty equivalent, whereas MVSK is consistently the best under the Sortino ratio. When T=12, MVSK and MV(1) are still the best and second best model

[^12]according to the Sharpe ratio, while the ranking is completely changed for Sortino and CEQ.
Book-to-market data show more stable rankings, with MVSK being the best model according to every measure for $\mathrm{T}=1$. Furthermore, the whole ranking is unchanged when short sales are allowed.in cases sich as $\mathrm{T}=60$ in Table $6 .{ }^{25}$

We assess ex-post performance also when investors have risk aversion coefficients equal to 2 and 10 (not reported here for sake of brevity). Generally, the relative performance of the models is not affected by the change in risk aversion. For instance, MV(1) is still the best model in International data both with risk aversion 2 and 10 when $\mathrm{T}=1$. What changes is the absolute value of the certainty equivalent, that is higher for low risk aversion coefficients, as the investor weighs less portfolio volatility.

One final remark concerns the effect of transaction costs when short sales are allowed. It is well-known that transaction costs are larger, reducing portfolio ex-post performances more with than without short sales, due to the large long and short positions entailed. Should this be the case, both MVSK and $1 / \mathrm{N}$ are likely to improve their relative performance with respect to MV models which yield less stable portfolio holdings.

## 5. Concluding comments

We find that the benchmark, equally weighted strategy never outperforms the ex-post performance of the optimizing ones in the three datasets under investigation.

We also uncover large ex-post gains from exploiting predictable moments up to the fourth order in international stock portfolios ranked according to Book-to-Market values. This result is the mirror image, cast in a portfolio choice setting, of previous finance literature highlighting the bad performance of Value portfolios in bear states. Importantly, regime-switching models deliver gains not only to an investor who cares about higher order moments, but also to an agent with mean-variance preferences.

The latter observation holds also when dealing with Industry portfolios, at least for short horizons. Descriptive statistics point to large co-kurtosis across Industry portfolios. Despite this evidence, mean variance strategies perform better than those considering predictability in higher order moments as well. This fact begs for an explanation. Guidolin and Nicodano (2008) indicate that third (fourth) moments yield considerable (little) additional welfare in sample over a dataset characterized by both types of nonnormalities. We conjecture that MVSK strategies perform poorly in the Industry dataset because these portfolio returns display large co-kurtosis - i.e. high volatility when other stocks are also highly volatile but little co-skewness. Further work may scrutinize whether the type of non-normality in the data affects ex-post gains from predicting higher order moments.

Our analysis also confirms previous results by Ang and Bekaert (2004): the out-of-sample cost of adopting mean-variance strategies is low when the investor diversifies across international equities, with no opportunity to shift into bonds in bear states. We see, however, that even in this dataset accounting for regimes and skewness, especially for long horizons, can improve risk-adjusted performances.

All datasets suggest that - to some extent - modelling the regime-switching nature of stock returns is

[^13]beneficial, but higher order moments matter only in two out of three datasets. This ought to be considered as a lower bound on the relevance of higher order moments for portfolio strategies, as it is conditional on the specific parametrization of our MVSK preferences. Their economic importance may be much larger for other types of preferences: for instance, allowing for investors' disappointment aversion, as in Hong et al. (2007), may boost gains from timing higher-order moments relative to the case of power utility.

In conclusion, allowing for a simple and parsimonious 2-state representation of the return distribution improves on ex-post performance always in the BM dataset, and for only a subset of investment horizons in the other two datasets. Predicting third and fourth moments, on top of the first two, need not always deliver gains even when descriptive statistics indicate the presence of sizeable non-normalities. The inclusion of ex-post transaction costs, which we leave for future work, may further increase the relative attractiveness of MVSK strategies because portfolio shares appear to be less sensitive to variations in expected returns.

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## 6. Appendix

## Moments of the wealth distribution

This appendix displays equations for moments of the wealth distribution, when there are autoregressive terms in the return process, and the number of regimes is set to $k=2$. Using (3) the $n$-th noncentral moment satisfies the recursions

$$
\begin{aligned}
M_{i, t+T}^{(n)}= & M_{i, t+T-1}^{(n)}(n) p_{i i} \exp \left(n \mu_{i}+n \sum_{j=1}^{p} a_{j, i} E_{t}\left[r_{t+T-j}\right]+\frac{n^{2}}{2} \sigma_{i}^{2}\right)+ \\
& +M_{-i, t+T-1}^{(n)}(n)\left(1-p_{-i-i}\right) \exp \left(n \mu_{i}+n \sum_{j=1}^{p} a_{j, i} E_{t}\left[r_{t+T-j}\right]+\frac{n^{2}}{2} \sigma_{i}^{2}\right)
\end{aligned}
$$

or

$$
\begin{aligned}
M_{1, t+1}^{(n)} & =\tilde{\alpha}_{1}^{(n)} M_{1, t}^{(n)}+\tilde{\beta}_{1}^{(n)} M_{2, t}^{(n)} \\
M_{2, t+1}^{(n)} & =\tilde{\alpha}_{2}^{(n)} M_{1, t}+\tilde{\beta}_{2}^{(n)} M_{2, t}^{(n)}
\end{aligned}
$$

where now

$$
\begin{aligned}
& \tilde{\alpha}_{1}^{(n)}=p_{11} \exp \left(n \mu_{1}+n \sum_{j=1}^{p} a_{j, 1} E_{t}\left[r_{t+T-j}\right]+\frac{n^{2}}{2} \sigma_{1}^{2}\right) \\
& \tilde{\beta}_{1}^{(n)}=\left(1-p_{22}\right) \exp \left(n \mu_{1}+n \sum_{j=1}^{p} a_{j, 1} E_{t}\left[r_{t+T-j}\right]+\frac{n^{2}}{2} \sigma_{1}^{2}\right) \\
& \tilde{\alpha}_{2}^{(n)}=\left(1-p_{11}\right) \exp \left(n \mu_{2}+n \sum_{j=1}^{p} a_{j, 2} E_{t}\left[r_{t+T-j}\right]+\frac{n^{2}}{2} \sigma_{2}^{2}\right) \\
& \tilde{\beta}_{2}^{(n)}=p_{22} \exp \left(n \mu_{2}+n \sum_{j=1}^{p} a_{j, 2} E_{t}\left[r_{t+T-j}\right]+\frac{n^{2}}{2} \sigma_{2}^{2}\right) .
\end{aligned}
$$

Subject to these changes, the earlier methods can be used with the only difference that terms such as $\exp \left(n \mu_{i}+\frac{n^{2}}{2} \sigma_{i}^{2}\right)$ have to be replaced by

$$
\exp \left(n \mu_{1}+n \sum_{j=1}^{p} a_{j, i} E_{t}\left[r_{t+T-j}\right]+\frac{n^{2}}{2} \sigma_{1}^{2}\right) .
$$

The term $\sum_{j=1}^{p} a_{j, i} E_{t}\left[r_{t+T-j}\right]$ may be decomposed in the following way:

$$
\sum_{j=1}^{p} a_{j, i} E_{t}\left[r_{t+T-j}\right]=I_{\{j>T\}} \sum_{j=1}^{p}\left(I_{\{j \geq T\}} a_{j, i} r_{t+T-j}+I_{\{j<T\}} a_{j, i} E_{t}\left[r_{t+T-j]}\right)\right.
$$

where $E_{t}\left[r_{t+1}\right], \ldots, E_{t}\left[r_{t+T-1}\right]$ can be evaluated recursively, c.f. Doan et al. (1984):

$$
\begin{aligned}
E_{t}\left[r_{t+1}\right]= & \pi_{1 t}\left(\mu_{1}+\sum_{j=1}^{p} a_{j, 1} r_{t-j}\right)+\left(1-\pi_{1 t}\right)\left(\mu_{2}+\sum_{j=1}^{p} a_{j, 2} r_{t-j}\right) \\
E_{t}\left[r_{t+2}\right]= & \boldsymbol{\pi}_{t}^{\prime} \mathbf{P e}_{1}\left(\mu_{1}+\sum_{j=1}^{p} a_{j, 1} E_{t}\left[r_{t+1}\right]\right)+\left(1-\boldsymbol{\pi}_{t}^{\prime} \mathbf{P e}_{1}\right)\left(\mu_{2}+\sum_{j=1}^{p} a_{j, 2} E_{t}\left[r_{t+1}\right]\right) \\
& \cdots \\
E_{t}\left[r_{t+T-1}\right]= & \boldsymbol{\pi}_{t}^{\prime} \mathbf{P}^{T-1} \mathbf{e}_{1}\left(\mu_{1}+\sum_{j=1}^{p} a_{j, 1} E_{t}\left[r_{t+T-2}\right]\right)+\left(1-\boldsymbol{\pi}_{t}^{\prime} \mathbf{P}^{T-1} \mathbf{e}_{1}\right)\left(\mu_{2}+\sum_{j=1}^{p} a_{j, 2} E_{t}\left[r_{t+T-2}\right]\right) .
\end{aligned}
$$

## Table 0 <br> List of asset-allocation models considered

This table lists the asset-allocation models we consider. The last column of the table gives the abbreviation used to refer to the strategy in the tables where we compare the performance of the optimal portfolio strategies. We consider six different time horizons, listed in the second column and the levels of relative risk aversions in the third column.
Models Horizons Risk Aversions Abbreviation
$\mathrm{T}=1,3,6,12,24,60,120 \quad 2,5,10$
Naive
1 Equal weighted
No predictability, no higher moments
2 Mean-variance (associated with MSIA(1,0))
3 Mean-variance, no shortsales
Predictability and higher moments preferences
42 Moment Pref with $\operatorname{MSIH}(2,0)$ returns
52 Moment Pref with $\operatorname{MSIH}(2,0)$ returns, no shortsales
Predictability and higher moments preferences
63 Moment Pref with $\operatorname{MSIH}(2,0)$ returns
73 Moment Pref with $\operatorname{MSIH}(2,0)$, no shortsales
83 Moment Pref with $\operatorname{MSIH}(2,0)$ returns

## Table 1

## Summary Statistics for Equity Returns

The table reports basic moments for monthly equity total return series for international portfolios from January 1988 to July 2008 (Panel A), Industries indices from July 1926 to July 2008 (Panel B) and International Book-to-Market portfolios from January 1975 to December 2007(Panel C) in the upper part of each panel. All returns are expressed in local currencies. Means, Median and Standard Deviations are annualized. The column Jarque-Bera reports the value of the Jarque-Bera statistics for normality, while $\mathrm{LB}(12)$ reports the $12^{\text {th }}$-order Ljung-Box statistic. The middle part of each panel reports the correlation and cokurtosis matrices, the lower part the co-skewness matrix. In the co-skewness matrix, coefficients above the main diagonal refer to the sample covariance between the square of the returns of the row portfolio and the level of returns of the column portfolio; coefficients below the main diagonal refer to the sample covariance between the level of the returns of the column portfolio/index and the square of returns of the row portfolio/index. In the correlation/co-kurtosis matrix, correlations are reported above the main diagonal and sample covariances between squared portfolio returns appear below the main diagonal.The symbols ${ }^{* *}$ and $*$ respectively denote statistical significance at $1 \%$ and $5 \%$.

|  | Mean | St. Dev. | Sharpe ratio | Median | Min. | Max. | Skewness | Kurtosis | Jarque-Bera | LB(12) | LB(12)-squares |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Pacific ex-Japan | 0.746* | 5.428** | 0.072 | 0.927 | -23.1 | 15.3 | -0.530* | 4.685* | 40.94** | 18.62 | 28.59** |
| Japan | -0.051 | 6.310** | -0.064 | -0.272 | -21.6 | 21.7 | 0.101 | 3.696 | 5.43 | 10.94 | 52.51* |
| Europe ex-UK | 0.766* | 4.928** | 0.084 | 1.103 | -15.6 | 13.8 | -0.542* | 4.059 | 23.73** | 15.48 | 25.22* |
| United Kingdom | 0.707 | 4.397** | 0.080 | 0.673 | -10.9 | 14.1 | 0.038 | 3.178 | 0.389 | 11.23 | 57.88** |
| North America | 0.756** | 3.924** | 0.103 | 1.093 | -14.3 | 10.4 | -0.441* | 3.714 | 13.309** | 7.77 | 34.88** |
| EM Latin America | 1.906** | 8.933** | 0.174 | 2.680 | -35.4 | 27.3 | -0.594* | 4.536* | 38.985** | 10.97 | 20.72 |
| EM Asia | 0.774 | 7.111** | 0.059 | 1.078 | -19.7 | 22.1 | -0.181 | 3.717 | 6.657* | 33.06** | 45.27** |
| EM Europe \& Middle East | 1.846** | 7.747** | 0.193 | 2.450 | -29.0 | 38.8 | 0.272 | 5.888** | 89.256** | 14.75 | 5.40 |
| Correlation and (variance) co-kurtosis matrices |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  | Pacificex-JP | Japan | EU ex-UK | UK | North Amr. | EM Latin Amr. | EM Asia | EM EU \& Middle East |  |
|  | Pacific ex-JP |  |  | 0.444** | 0.592** | 0.621** | 0.601** | 0.545** | 0.785** | 0.424** |  |
|  | Japan |  | 1.292 |  | 0.462** | 0.480** | 0.368** | 0.321** | 0.406** | 0.221** |  |
|  | EU ex-UK |  | 1.894 | 1.715 |  | 0.744** | 0.669** | 0.410** | 0.508** | 0.467** |  |
|  | UK |  | 1.533 | 1.459 | 2.317* |  | 0.664** | 0.394** | 0.442** | 0.353* |  |
|  | North Amr. |  | 1.766 | 1.346 | 2.801* | 2.138** |  | 0.500** | 0.551** | 0.404** |  |
|  | EM Latin Amr. |  | 1.923 | 1.806 | 2.012* | 1.379 | 2.255 |  | 0.491** | 0.479** |  |
|  | EM Asia |  | 3.149** | 1.440 | 2.114 | 1.488 | 1.979 | 1.826 |  | 0.475** |  |
|  | EM EU \& Middle | East | 1.407 | 1.503 | 2.000 | 1.384 | 2.250 | 2.286 | 1.586 |  |  |
| Co-skewness matrix |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  | Pacificex-JP | Japan | EU ex-UK | UK | North Amr. | EM Latin Amr. | EM Asia | EM EU \& Middle East |  |
|  | Pacific ex-JP |  |  | -0.273 | -0.360 | -0.203 | -0.352* | -0.385 | -0.364 | -0.299* |  |
|  | Japan |  | -0.159 |  | -0.052 | 0.071 | -0.078 | -0.307 | -0.160 | -0.115 |  |
|  | EU ex-UK |  | -0.374** | -0.306 |  | -0.351 | -0.496* | -0.354 | -0.470** | -0.344 |  |
|  | UK |  | -0.043 | -0.090 | -0.178 |  | -0.080 | -0.117 | -0.118 | -0.187 |  |
|  | North Amr. |  | -0.393** | -0.231 | -0.520** | -0.238 |  | -0.475* | -0.410* | -0.419 |  |
|  | EM Latin Amr. |  | -0.400** | -0.420* | -0.213 | -0.271 | -0.380 |  | -0.269 | -0.337 |  |
|  | EM Asia |  | -0.292 | -0.263 | -0.372* | -0.209 | -0.273 | -0.277 |  | -0.288* |  |
|  | EM EU \& Middle | East | -0.220 | -0.131 | -0.126 | -0.218 | -0.344 | -0.565 | -0.260 |  |  |

Table 1 (cont'ed)
Summary Statistics for Equity Returns

|  | Mean | St. Dev. | Sharpe ratio | Median | Min. | Max. | Skewness | Kurtosis | Jarque-Bera | LB(12) | LB(12)-squares |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Non Durables | 0.978** | 4.691** | 0.143 | 1.090 | -24.5 | 34.4 | -0.029 | 8.843** | 1401.5** | 36.17** | 333.9** |
| Durables | 1.074** | 7.593** | 0.101 | 1.000 | -34.8 | 79.7 | 1.203 | 18.42* | 9993.4** | 51.38** | 266.4** |
| Manufacturing | 1.034** | 6.322** | 0.115 | 1.330 | -29.8 | 57.4 | 0.978 | 15.59** | 6761.2** | 39.407** | 440.3** |
| Energy | 1.097** | 5.983** | 0.132 | 0.860 | -26.0 | 33.5 | 0.238 | 6.183** | 425.02** | 23.27* | 251.4** |
| Hi Tech | 1.094** | 7.437** | 0.106 | 1.220 | -33.8 | 53.4 | 0.296 | 9.030** | 1506.7** | 26.91** | 562.3** |
| Telecommunications | 0.831** | 4.594** | 0.115 | 0.880 | -21.6 | 28.2 | 0.056 | 6.277** | 441.18** | 29.49** | 342.2** |
| Shops/Distribution | 0.975** | 5.884** | 0.114 | 1.130 | -30.2 | 37.1 | -0.016 | 8.501** | 1242.0** | 56.72** | 458.6** |
| Health | 1.089** | 5.766** | 0.136 | 1.070 | -34.7 | 38.7 | 0.171 | 10.210** | 2136.2** | 53.59** | 605.9** |
| Utilities | 0.902** | 5.685** | 0.105 | 1.050 | -33.0 | 43.2 | 0.095 | 10.61** | 2379.7** | 52.12** | 622.3** |
| Other | 0.921** | 6.473** | 0.095 | 1.260 | -30.0 | 58.7 | 0.971 | 16.83** | 8006.3** | 64.55** | 490.0** |
| Correlation and (variance) co-kurtosis matrices |  |  |  |  |  |  |  |  |  |  |  |
|  |  | Non Durables | Durables | Manufactur <br> e | Energy | Hi Tech | Telecom | Shops | Health | Utilities | Other |
| Non Durables |  |  | 0.754** | 0.851** | 0.616** | 0.736** | 0.671** | 0.866** | 0.801** | 0.707** | 0.847** |
| Durables |  | 10.490* |  | 0.873** | 0.607** | 0.779** | 0.618** | 0.798** | 0.649** | 0.635** | 0.802** |
| Manufacturing |  | 10.150** | 15.873* |  | 0.723** | 0.862** | 0.671** | 0.841** | 0.762** | 0.703** | 0.905** |
| Energy |  | 5.297** | 7.095** | 7.328** |  | 0.607** | 0.495** | 0.576** | 0.562** | 0.617** | 0.689** |
| Hi Tech |  | 7.546** | 11.109* | 10.872** | 5.273** |  | 0.676** | 0.785** | 0.723** | 0.624** | 0.798** |
| Telecommunications |  | 4.614** | 5.257** | 5.930** | 3.265** | 5.248** |  | 0.670** | 0.600** | 0.635** | 0.695** |
| Shops/Distribution |  | 7.880** | 9.697** | 9.688** | 4.810** | 7.499** | 4.878** |  | 0.740** | 0.655** | 0.824** |
| Health |  | 8.130** | 9.952** | 10.163** | 4.987** | 7.842** | 5.079** | 7.860** |  | 0.625** | 0.741** |
| Utilities |  | 6.660** | 8.297** | 9.188** | 4.686** | 7.514** | 5.839** | 7.502** | 7.707** |  | 0.740** |
| Other |  | 8.496** | 10.936** | 13.017** | 6.500** | 9.662** | 6.876* | 8.927** | 9.863** | 10.436** |  |
|  | Co-Skewness matrices |  |  |  |  |  |  |  |  |  |  |
|  |  | Non Durables | Durables | $\begin{gathered} \hline \text { Manufactur } \\ \text { e } \\ \hline \end{gathered}$ | Energy | Hi Tech | Telecom | Shops | Health | Utilities | Other |
| Non Durables |  |  | 0.248 | 0.196 | -0.032 | 0.058 | -0.153 | -0.054 | -0.021 | -0.077 | 0.046 |
| Durables |  | 0.633 |  | 1.042 | 0.554 | 0.754 | 0.164 | 0.577 | 0.615 | 0.397 | 0.663 |
| Manufacturing |  | 0.505 | 0.981 |  | 0.459 | 0.684 | 0.259 | 0.477 | 0.600 | 0.472 | 0.777 |
| Energy |  | -0.013 | 0.248 | 0.239 |  | 0.128 | -0.052 | -0.064 | 0.118 | 0.029 | 0.180 |
| Hi Tech |  | 0.172 | 0.451 | 0.460 | 0.156 |  | -0.017 | 0.167 | 0.234 | 0.155 | 0.333 |
| Telecommunications |  | -0.093 | -0.085 | 0.043 | -0.126 | -0.030 |  | -0.096 | -0.041 | 0.064 | 0.109 |
| Shops/Distribution |  | -0.051 | 0.228 | 0.187 | -0.058 | 0.069 | -0.104 |  | -0.051 | -0.024 | 0.119 |
| Health |  | 0.061 | 0.325 | 0.363 | 0.150 | 0.192 | 0.030 | 0.037 |  | 0.065 | 0.311 |
| Utilities |  | -0.046 | 0.120 | 0.228 | -0.032 | 0.098 | 0.089 | -0.020 | 0.009 |  | 0.302 |
| Other |  | 0.337 | 0.658 | 0.802 | 0.380 | 0.555 | 0.370 | -0.020 | 0.555 | 0.579 |  |

Table 1 (cont'ed)
Summary Statistics for Equity Returns
Panel C (International Book-to-Market Sorted Portfolio Local Returns, 1975:01-2007:12)

|  | Mean | St. Dev. | Sharpe ratio | Median | Min. | Max. | Skewness | Kurtosis | Jarque-Bera | LB(12) | LB(12)-squares |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| World | 0.694** | 3.844** | 0.055 | 1.000 | -22.0 | 12.8 | -0.985** | 6.832 | 306.4** | 13.88 | 6.851 |
| EU ex-UK ex-Scand Value | 1.326** | 4.740** | 0.178 | 1.640 | -18.7 | 16.4 | -0.486* | 5.130* | 90.47** | 24.06* | 22.72* |
| EU ex-UK ex-Scand Growth | 0.967** | 4.294** | 0.112 | 1.280 | -24.9 | 14.9 | -0.733 | 6.915 | 288.4** | 18.78 | 15.89 |
| United Kingdom Value | 1.725** | 6.187** | 0.201 | 1.700 | -27.0 | 45.5 | 0.845 | 10.95 | 1088.9** | 13.80 | 44.02** |
| United Kingdom Growth | 1.353** | 5.851** | 0.148 | 1.325 | -27.9 | 53.8 | 1.610* | 20.44* | 5186.8** | 12.60 | 11.83 |
| Asia \& Pacific Value | 1.300** | 5.123** | 0.159 | 1.000 | -25.0 | 19.1 | -0.095 | 5.615* | 113.4** | 11.57 | 56.82** |
| Asia \& Pacific Growth | 0.401 | 4.936** | -0.017 | 0.525 | -18.4 | 25.1 | -0.003 | 5.214 | 80.88** | 11.21 | 81.93** |
| Scandinavia Value | 1.676** | 6.428** | 0.185 | 1.765 | -22.1 | 25.8 | 0.175 | 4.358* | 32.46** | 29.48** | 25.14* |
| Scandinavia Growth | 1.486** | 6.270** | 0.160 | 1.770 | -21.4 | 25.5 | 0.037 | 4.763** | 51.36** | 22.75* | 80.34** |
| United States Value | 1.081** | 4.769** | 0.125 | 1.255 | -24.3 | 14.2 | -0.473 | 4.901 | 74.41** | 7.301 | 11.83 |
| United States Growth | 1.448** | 4.302** | 0.224 | 1.660 | -20.4 | 23.7 | -0.177 | 7.616** | 353.7** | 19.06 | 29.92** |


| Correlation and (variance) co-kurtosis matrices |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | World | EU ex-UK exScand Value | Uex-UK ex Scand Growth | UK Value | UK Growth | Asia Pacific Value | Asia Pacific Growth | Scandinavi a Value | Scandinavia Growth | United <br> States Value | United States Growth |
| World |  | 0.742** | 0.792** | 0.632** | 0.644** | 0.593** | 0.681** | 0.525** | 0.625** | 0.859** | 0.804** |
| EU ex-UK ex-Scand Value | 4.624** |  | 0.850** | 0.573** | 0.506** | 0.470** | 0.441** | 0.584** | 0.542** | 0.532** | 0.627** |
| EU ex-UK ex-Scand Growth | 6.160** | 4.682** |  | 0.550** | 0.584** | 0.434** | 0.513** | 0.499** | 0.619** | 0.621** | 0.593** |
| United Kingdom Value | 4.901** | 3.204** | 4.632** |  | 0.786** | 0.403** | 0.389** | 0.419** | 0.377** | 0.464** | 0.556** |
| United Kingdom Growth | 5.620** | 3.397** | 5.560* | 13.637* |  | 0.317** | 0.357** | 0.370** | 0.407** | 0.546** | 0.563** |
| Asia \& Pacific Value | 4.452** | 3.157** | 4.368* | 3.107** | 2.962* |  | 0.649** | 0.412** | 0.332** | 0.337** | 0.360** |
| Asia \& Pacific Growth | 2.995** | 2.047** | 2.811** | 1.837** | 1.806** | 3.071** |  | 0.344** | 0.454** | 0.409** | 0.350** |
| Scandinavia Value | 2.827** | 2.515** | 2.820** | 2.149** | 2.162** | 2.325** | 1.620** |  | 0.643** | 0.372** | 0.432** |
| Scandinavia Growth | 3.330** | 2.453** | 3.569** | 2.132** | 2.200** | 2.267** | 1.989** | 2.160** |  | 0.546** | 0.434** |
| United States Value | 5.050** | 3.166** | 4.520* | 3.428** | 3.819* | 3.078 | 1.873** | 2.223** | 2.789** |  | 0.784** |
| United States Growth | 5.582** | 3.727** | 4.887* | 6.251* | 8.375* | 3.082* | 1.748** | 2.256** | 2.219** | 4.460** |  |
| Co-Skewness matrices |  |  |  |  |  |  |  |  |  |  |  |
|  | World | EU ex-UK exScand Value | Jex-UK ex Scand Growth | UK Value | UK Growth | Asia Pacific Value | Asia Pacific Growth | Scandinavi a Value | Scandinavia Growth | United States Value | United States Growth |
| World |  | -0.837* | -0.976* | -0.516 | -0.499 | -0.829* | -0.572* | -0.650* | -0.681* | -0.797 | -0.733 |
| EU ex-UK ex-Scand Value | -0.661* |  | -0.595* | -0.359 | -0.416 | -0.501* | -0.335* | -0.437* | -0.433* | -0.564* | -0.552* |
| EU ex-UK ex-Scand Growth | -0.881 | -0.685 |  | -0.536 | -0.499 | -0.771 | -0.445 | -0.642* | -0.552 | -0.783 | -0.741 |
| United Kingdom Value | -0.040 | -0.116 | -0.119 |  | 0.994 | -0.129 | -0.049 | -0.249 | -0.208 | -0.097 | 0.340 |
| United Kingdom Growth | -0.138 | -0.028 | 0.048 | 1.236 |  | -0.199 | -0.041 | -0.168 | -0.056 | 0.015 | 0.589 |
| Asia \& Pacific Value | -0.583 | -0.410 | -0.528 | -0.278 | -0.404 |  | -0.239 | -0.436 | -0.437 | -0.524 | -0.533 |
| Asia \& Pacific Growth | -0.301 | -0.205 | -0.197 | -0.097 | -0.148 | -0.154 |  | -0.329** | -0.307* | -0.259* | -0.285* |
| Scandinavia Value | -0.339 | -0.317 | -0.366 | -0.191 | -0.280 | -0.216 | -0.123 |  | -0.109 | -0.309 | -0.300 |
| Scandinavia Growth | -0.389 | -0.349* | -0.353 | -0.308 | -0.231 | -0.399* | -0.243 | -0.134 |  | -0.225 | -0.324 |
| United States Value | -0.644 | -0.635** | -0.790* | -0.376 | -0.407 | -0.583 | -0.360* | -0.437 | -0.424 |  | -0.451 |
| United States Growth | -0.501 | -0.559 | -0.668 | 0.017 | 0.006 | -0.489 | -0.343 | -0.424 | -0.412 | -0.415 |  |

** Statistical significance at $1 \%$; * Statistical significance at 5\%.

## Table 2

## Model Selection Statistics

The table reports estimates for the multivariate Markov switching conditionally heteroskedastic VAR model:

$$
\mathrm{r}_{t}=\mu_{s_{t}}+\sum_{j=1}^{p} \mathrm{~A}_{j, s_{t}} \mathrm{r}_{t-j}+\varepsilon_{t}
$$

where $\mu_{s t}$ is the intercept vector in state $\mathrm{s}_{t}, \mathrm{~A}_{j, s_{t}}$ is the matrix of autoregressive coefficients associated with leg $j \geq 1$ in state $\mathrm{s}_{t}$ and $\varepsilon_{t}=\left(\varepsilon_{1 t}, \ldots \varepsilon_{h t}\right){ }^{\prime} \sim \mathrm{N}\left(\mathbf{0}, \Omega_{s_{t}}\right)$. The unobserved state variable $\mathrm{s}_{t}$ Is governed by a first-order Markov chain that can assume k distinct values. P autoregressive terms are considered. The sample period is 1988:01-2008:08 for Panel A (International porfolios), 1926:07-2008:07 for Panel B (Industries) and 1975:01-2007:12 for Panel C (Book-to-Market). MISIAH( $k, p$ ) stands for Markov Switching Intercept Autoregressive Heteroskedasticity Model with $k$ states and $p$ autoregressive lags.

Panel A (International MSCI USD Returns, 1988:01-2008:08)

| Model (K,p) | Log-likelihood LR Statistic | Davies' approx. pr <br> value | BIC | HQ | AIC | Number of <br> parameters | Number of obs. Saturation ratio |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Single-state models | International MSCI USD Returns, 1988:01-2008:07 |  |  |  |  |  |

Table 2 (cont'ed)
Model Selection Statistics

| Model ( $\mathrm{K}, \mathrm{p}$ ) | Log-likelihood | LR Statistic | Davies' approx. pvalue | BIC | HQ | AIC | Number of parameters | Number of obs. | Saturation ratio | Tests |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| CRSP Industry Returns, 1926:07-2008:07 |  |  |  |  |  |  |  |  |  |  |
| Single-state models |  |  |  |  |  |  |  |  |  |  |
| MSIA $(1,0)$ | 19354.00 | - |  | -38.843 | -39.043 | -39.166 | 65 | 9850 | 151.5 |  |
| MSIA $(1,1)$ | 19481.34 | - |  | -38.441 | -38.949 | -39.261 | 165 | 9850 | 59.7 |  |
| Two-state models |  |  |  |  |  |  |  |  |  |  |
| MSI( 2,0 ) | 19428.38 | 148.767 | 0.000 | -38.910 | -39.147 | -39.292 | 77 | 9850 | 127.9 | H: 2028.35 (0.000) |
| MSIH $(2,0)$ | 20442.56 | 2177.118 | 0.000 | -40.391 | -40.797 | -41.047 | 132 | 9850 | 74.6 |  |
| MSH $(2,0)$ | 20242.57 | 1777.147 | 0.000 | -40.344 | -40.720 | -40.950 | 122 | 9850 | 80.7 | I: 399.97 (0.000) |
| MSIA $(2,1)$ | 19943.31 | 923.937 | 0.000 | -38.595 | -39.448 | -39.972 | 277 | 9840 | 35.5 |  |
| $\operatorname{MSIAH}(2,1)$ | 20514.06 | 2065.444 | 0.000 | -39.370 | -40.393 | -41.020 | 332 | 9840 | 29.6 |  |
| Three-state models |  |  |  |  |  |  |  |  |  |  |
| MSI( 3,0 ) | 19532.68 | 357.374 | 0.000 | -39.024 | -39.476 | -39.304 | 91 | 9850 | 108.2 |  |
| MSIH $(3,0)$ | 20504.22 | 2300.450 | 0.000 | -40.226 | -40.845 | -41.225 | 201 | 9850 | 49.0 |  |
| $\operatorname{MSIA}(3,1)$ | 20106.58 | 1250.476 | 0.000 | -38.129 | -39.333 | -40.072 | 391 | 9840 | 25.2 |  |
| MSIAH $(3,1)$ | 20704.65 | 2446.620 | 0.000 | -38.574 | -40.117 | -41.064 | 501 | 9840 | 19.6 |  |
| Four-state models |  |  |  |  |  |  |  |  |  |  |
| MSI(4,0) | 19607.79 | 507.578 | 0.000 | -39.064 | -39.393 | -39.596 | 107 | 9850 | 92.1 |  |
| MSIH $(4,0)$ | 20423.45 | 2138.898 | 0.000 | -40.136 | -40.973 | -41.487 | 272 | 9850 | 36.2 |  |
| $\mathrm{MSIH}(4,0)-\mathrm{VAR}(1)$ | 20826.51 | 2690.348 | 0.000 | -39.725 | -40.871 | -41.574 | 372 | 9840 | 26.5 |  |

Table 2 (cont'ed)
Model Selection Statistics

Panel C (International Book-to-Market Sorted Portfolio Local Returns, 1975:01-2007:12)

| Model (K,p) | Log-likelihood | LR Statistic | Davies' approx. $p$. value | BIC | HQ | AIC | Number of parameters | Number of obs. | Saturation ratio | Tests |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| International Book-to-Market Sorted Portfolio Local Returns, 1975:01-2007:12 |  |  |  |  |  |  |  |  |  |  |
| Single-state models |  |  |  |  |  |  |  |  |  |  |
| MSIA 1,0 ) | 8772.32 | - |  | -43.142 | -43.609 | -43.916 | 77 | 4356 | 56.6 |  |
| MSIA $(1,1)$ | 8929.72 |  |  | -41.171 | -43.190 | -44.516 | 198 | 4345 | 21.9 |  |
| Two-state models |  |  |  |  |  |  |  |  |  |  |
| MSI( 2,0 ) | 8830.76 | 116.877 | 0.000 | -43.240 | -43.787 | -44.145 | 90 | 4356 | 48.4 | H: 505.68 (0.000) |
| MSIH $(2,0)$ | 9083.60 | 622.555 | 0.000 | -43.521 | -44.468 | -45.089 | 156 | 4356 | 27.9 |  |
| MSH $(2,0)$ | 9038.61 | 532.590 | 0.000 | -43.513 | -44.389 | -44.965 | 145 | 4356 | 30.0 | I: 89.98 (0.000) |
| MSIA $(2,1)$ | 9123.80 | 388.169 | 0.000 | -41.171 | -43.190 | -44.516 | 332 | 4345 | 13.1 |  |
| MSIAH $(2,1)$ | 9240.63 | 621.826 | 0.000 | -40.764 | -43.184 | -44.773 | 398 | 4345 | 10.9 | H: 233.66 (0.000) |
| Three-state models |  |  |  |  |  |  |  |  |  |  |
| MSI( 3,0 ) | 8854.06 | 163.463 | 0.000 | -43.132 | -43.769 | -44.187 | 105 | 4356 | 41.5 |  |
| MSIH $(3,0)$ | 9212.01 | 879.363 | 0.000 | -42.946 | -44.384 | -45.328 | 237 | 4356 | 18.4 |  |
| MSIA $(3,1)$ | 9344.81 | 830.185 | 0.000 | -40.232 | -43.078 | -44.946 | 468 | 4345 | 9.3 |  |
| MSIAH $(3,1)$ |  | No converge | achieved (too many | paramet |  |  | 600 | 4345 | 7.2 |  |
| Four-state models |  |  |  |  |  |  |  |  |  |  |
| MSI(4,0) | 8894.16 | 243.666 | 0.000 | -43.077 | -43.818 | -44.304 | 122 | 4356 | 35.7 |  |
| MSIH(4,0) | 9375.96 | 1207.270 | 0.000 | -42.520 | -44.463 | -45.737 | 320 | 4356 | 13.6 |  |
| $\underline{\operatorname{MSIH}(4,0)-\operatorname{VAR}(1)}$ | 9452.93 | 1046.411 | 0.000 | -41.188 | -43.870 | -45.630 | 441 | 4345 | 9.9 |  |

## Table 3

## Estimated Markov Switching Models

The table shows estimation results for the regime switching model

$$
\mathrm{r}_{t}=\mu_{s_{t}}+\varepsilon_{t}
$$

$\mathrm{r}_{t}$ is the vector collecting monthly total return series, $\mu_{s_{t}}$ is the intercept vector in state $\mathrm{s}_{t}$ and $\varepsilon_{t}=\left(\varepsilon_{1 t}, \ldots \varepsilon_{h t}\right)^{\prime} \sim$ $\mathrm{N}\left(\mathbf{0}, \Omega_{s_{t}}\right)$. Panel A shows the estimation results for the International portfolios dataset, Panel B for the Industries one, Panel C for the Book-to-Market portfolios one. Each Panel reports the results for the single-state model, k=1, (Panel A), for the two-state model $\mathrm{k}=2$ (Panel B) and its Markov chain properties (Panel C) (ergodic probabilites and average state duration).

Panel A (International MSCI USD Returns, 1988:01-2008:08)

|  | Panel A - SINGLE STATE MODEL |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Pacific EX <br> JP | Japan | Europe EX UK | UK | North America | EM Latin America | EM Asia | EM Europe \& Middle East |
| 1. Mean returns | 0.746* | -0.051 | 0.766* | 0.707* | 0.756** | 1.906** | 0.774 | 1.846** |
| 2. Correlations/Volatilities |  |  |  |  |  |  |  |  |
| Pacific EX JP | 5.428** |  |  |  |  |  |  |  |
| JP | 0.444** | 6.310** |  |  |  |  |  |  |
| Europe EX UK | 0.592** | 0.462** | 4.928** |  |  |  |  |  |
| UK | 0.621** | 0.480** | 0.744** | 4.397** |  |  |  |  |
| North America | 0.601** | 0.368** | 0.669** | 0.664** | 3.924** |  |  |  |
| EM Latin America | 0.545** | 0.321** | 0.410** | 0.394** | 0.500** | 8.932** |  |  |
| EM Asia | 0.785** | 0.406** | 0.508** | 0.442** | 0.551** | 0.491** | 7.111** |  |
| EM Europe and Middle East | 0.424** | 0.221** | 0.467** | 0.352** | 0.404** | 0.479** | 0.475** | 7.747** |
|  | Panel B - TWO-STATE MODEL |  |  |  |  |  |  |  |
|  | Pacific EX JP | JP | Europe EX UK | UK | North America | EM Latin America | EM Asia | EM Europe \& Middle East |
| 1. Mean returns |  |  |  |  |  |  |  |  |
| Bear/High Volatility State | -0.386* | -1.325** | -0.087 | -0.051 | 0.384 | 0.496 | 0.336 | 0.227 |
| Bull/Low Volatility State | 1.371** | 0.653 | 1.238** | 1.125** | 0.961** | 2.685** | 1.017* | 2.740** |
| 2. Correlations/Volatilities |  |  |  |  |  |  |  |  |
| Bear/High Volatility State |  |  |  |  |  |  |  |  |
| Pacific EX JP | 7.081** |  |  |  |  |  |  |  |
| JP | 0.399** | 7.524** |  |  |  |  |  |  |
| Europe EX UK | 0.498** | 0.473** | 5.850** |  |  |  |  |  |
| UK | 0.554** | 0.514** | 0.772** | 4.478** |  |  |  |  |
| North America | 0.561** | 0.355** | 0.621** | 0.509** | 4.284** |  |  |  |
| EM Latin America | 0.514** | 0.336* | 0.234* | 0.365* | 0.475** | 1.159** |  |  |
| EM Asia | 0.779** | 0.365** | 0.435** | 0.364** | 0.550** | 0.478** | 9.173** |  |
| EM Europe and Middle East | 0.333* | 0.002 | 0.375** | 0.448** | 0.483** | 0.407** | 0.411** | 8.406** |
| Bull/Low Volatility State |  |  |  |  |  |  |  |  |
| Pacific EX JP 4.095** | 4.095** |  |  |  |  |  |  |  |
| JP | 0.478** 5.376** |  |  |  |  |  |  |  |
| Europe EX UK |  |  |  |  |  |  |  |  |
| UK $\quad 0.712^{* *} 00.443^{* *} 00.727^{* *} 0.281^{* *}$ |  |  |  |  |  |  |  |  |
| North America $\quad 0.665^{* *}$ 0.372** $0.713^{* *} 00.767^{* *} 03.681^{* *}$ |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |
| EM Europe and Middle East $0.512 * *$ $0.386^{* *}$ $0.532 * *$ $0.267^{*}$ $0.335^{*}$ $0.554^{* *}$ $0.554^{* *}$ $7.176^{* *}$ |  |  |  |  |  |  |  |  |
| 3. Transition probabilities Bear/High Volatility State Bull/Low Volatility State |  |  |  |  |  |  |  |  |
| Bear/High Volatility State Bull/Low Volatility State | 0.807** |  |  |  | 0.193 |  |  |  |
|  | 0.103 |  |  |  | 0.897** |  |  |  |
| Panel C - MARKOV CHAIN PROPERTIES, TWO-STATE MODEL |  |  |  |  |  |  |  |  |
|  | Bear | Bull |  |  |  |  | Bear | Bull |
| Ergodic Probs | 0.348 | 0.652 |  | Average d | uration (in | months) | 5.18 | 9.70 |

Table 3 (cont'ed)
Estimated Markov Switching Models

Panel B (CRSP Industry Returns, 1926:07-2008:07)

|  | Panel A - SINGLE STATE MODEL |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | NonDurables | Durables | Manufacture | Energy | Hi Tech | Telecom | Shops/ Distrib. | Health | Utilities | Other |
| 1. Mean returns | 0.978** | 1.074** | 1.034** | 1.097** | 1.094** | 0.831** | 0.975** | 1.089** | 0.902** | 0.921** |
| 2. Correlations/Volatilities |  |  |  |  |  |  |  |  |  |  |
| Non-Durables | 4.691** |  |  |  |  |  |  |  |  |  |
| Durables | 0.754** | 7.593** |  |  |  |  |  |  |  |  |
| Manufacture | 0.851** | 0.873** | 6.322** |  |  |  |  |  |  |  |
| Energy | 0.616** | 0.607** | 0.723** | 5.983** |  |  |  |  |  |  |
| Hi Tech | 0.735** | 0.779** | 0.862** | 0.609** | 7.437** |  |  |  |  |  |
| Telecom | 0.671** | 0.618** | 0.671** | 0.495** | 0.676** | 4.594** |  |  |  |  |
| Shops/ Distrib. | 0.866** | 0.798** | 0.841** | 0.576** | 0.785** | 0.670** | 5.884** |  |  |  |
| Health | 0.801** | 0.649** | 0.762** | 0.562** | 0.723** | 0.600** | 0.740** | 5.766** |  |  |
| Utilities | 0.707** | 0.635** | 0.703** | 0.617** | 0.624** | 0.635** | 0.655** | 0.625** | 5.685** |  |
| Other | 0.847** | 0.802** | 0.905** | 0.689** | 0.798** | 0.695** | 0.824** | 0.741** | 0.740** | 6.473** |
| Panel B - TWO STATE MODEL MSIH(2,0) |  |  |  |  |  |  |  |  |  |  |
|  | NonDurables | Durables | Manufacture | Energy | Hi Tech | Telecom | Shops/ Distrib. | Health | Utilities | Other |
| 1. Mean returns |  |  |  |  |  |  |  |  |  |  |
| Bear/High Volatility State | 0.162 | 0.751 | 0.718 | 0.760 | 0.650 | 0.238 | 0.081 | 0.483 | 0.447 | 0.111 |
| Bull/Low Volatility State | 1.212** | 1.167** | 1.125** | 1.194** | 1.221** | 1.001** | 1.232** | 1.264** | 1.032** | 1.154** |
| 2. Correlations/Volatilities |  |  |  |  |  |  |  |  |  |  |
| Bear/High Volatility State |  |  |  |  |  |  |  |  |  |  |
| NoDur | 7.233** |  |  |  |  |  |  |  |  |  |
| Durbl | 0.666** | 13.181** |  |  |  |  |  |  |  |  |
| Manuf | 0.810** | 0.810** | 10.848** |  |  |  |  |  |  |  |
| Enrgy | 0.473** | 0.492** | 0.624** | 9.318** |  |  |  |  |  |  |
| HiTec | 0.695** | 0.711** | 0.837** | 0.490** | 12.657** |  |  |  |  |  |
| Telcm | 0.653** | 0.511** | 0.590** | 0.393* | 0.488** | 7.395** |  |  |  |  |
| Shops | 0.832** | 0.718** | 0.796** | 0.417** | 0.732** | 0.591** | 9.543** |  |  |  |
| Hlth | 0.775** | 0.543** | 0.739** | 0.434** | 0.686** | 0.522** | 0.682** | 8.903** |  |  |
| Utils | 0.671** | 0.499** | 0.600** | 0.532** | 0.470** | 0.621** | 0.528** | 0.514** | 10.021** |  |
| Other | 0.832** | 0.746** | 0.890** | 0.605** | 0.770** | 0.625** | 0.801** | 0.706** | 0.662** | 10.987** |
| Bull/Low Volatility State |  |  |  |  |  |  |  |  |  |  |
| NoDur | 3.607** |  |  |  |  |  |  |  |  |  |
| Durbl | 0.826** | 4.913** |  |  |  |  |  |  |  |  |
| Manuf | 0.893** | 0.905** | 4.186** |  |  |  |  |  |  |  |
| Enrgy | 0.739** | 0.689** | 0.797** | 4.586** |  |  |  |  |  |  |
| HiTec | 0.772** | 0.814** | 0.875** | 0.691** | 4.999** |  |  |  |  |  |
| Telcm | 0.683** | 0.686** | 0.724** | 0.574** | 0.795** | 3.360** |  |  |  |  |
| Shops | 0.894** | 0.850** | 0.872** | 0.698** | 0.821** | 0.723** | 4.247** |  |  |  |
| Hlth | 0.823** | 0.729** | 0.787** | 0.670** | 0.756** | 0.660** | 0.785** | 4.453** |  |  |
| Utils | 0.744** | 0.698** | 0.754** | 0.681** | 0.700** | 0.649** | 0.729** | 0.706** | 3.550** |  |
| Other | 0.866** | 0.832** | 0.914** | 0.752** | 0.814** | 0.738** | 0.838** | 0.771** | 0.778** | 4.351** |
| 3. Transition probabilities |  | Bear/H | igh Volatility S | ate |  |  | Bull/ | w Volatilit | tate |  |
| Bear/High Volatility State |  |  | 0.822** |  |  |  |  | 0.178 |  |  |
| Bull/Low Volatility State |  |  | 0.051 |  |  |  |  | 0.949** |  |  |
| Panel C - MARKOV CHAIN PROPERTIES, TWO-STATE MODEL |  |  |  |  |  |  |  |  |  |  |
|  | Bear | Bull |  |  |  |  |  |  | Bear | Bull |
| Ergodic Probs | 0.224 | 0.776 |  |  |  | Average dur | uration (in | nths) | 5.60 | 19..46 |

Table 3 (cont'ed)
Estimated Markov Switching Models
Panel C (International Book-to-Market Sorted Portfolio Local Returns, 1975:01-2007:12)

|  | Panel A - SINGLE STATE MODEL |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | World | EU ex-UK exScand Value | EU ex-UK exScand Growth | UK Value | UK Growth | Asia \& Pacific Value | Asia \& Pacific Growth | Scandinavia Value | Scandinavia Growth | US Value | US Growth |
| 1. Mean returns | 0.694** | 1.326** | 0.967** | 1.725** | 1.353** | 1.300** | 0.401 | 1.676** | 1.486** | 1.081** | 1.448** |
| 2. Correlations/Volatilities |  |  |  |  |  |  |  |  |  |  |  |
| World | 3.844** |  |  |  |  |  |  |  |  |  |  |
| EU ex-UK ex-Scand Value | 0.742** | 4.740** |  |  |  |  |  |  |  |  |  |
| EU ex-UK ex-Scand Growth | 0.792** | 0.850** | 4.294** |  |  |  |  |  |  |  |  |
| United Kingdom Value | 0.632** | 0.573** | 0.549** | 6.187** |  |  |  |  |  |  |  |
| United Kingdom Growth | 0.644** | 0.506** | 0.584** | 0.786** | 5.851** |  |  |  |  |  |  |
| Asia \& Pacific Value | 0.593** | 0.470** | 0.434** | 0.403** | 0.317** | 5.123** |  |  |  |  |  |
| Asia \& Pacific Growth | 0.681** | 0.441** | 0.513** | 0.389** | 0.357** | 0.649** | 4.936** |  |  |  |  |
| Scandinavia Value | 0.525** | 0.584** | 0.500** | 0.419** | 0.370** | 0.412** | 0.344** | 6.428** |  |  |  |
| Scandinavia Growth | 0.625** | 0.542** | 0.619** | 0.377** | 0.407** | 0.332** | 0.454** | 0.643** | 6.270** |  |  |
| United States Value | 0.859** | 0.532** | 0.621** | 0.464** | 0.550** | 0.337** | 0.409** | 0.372** | 0.546** | 4.769** |  |
| United States Growth | 0.804** | 0.623** | 0.593** | 0.556** | 0.563** | 0.360** | 0.350** | 0.432** | 0.434** | 0.784** | 4.302** |
|  | Panel B - TWO STATE MODEL MSIH( 2,0 ) |  |  |  |  |  |  |  |  |  |  |
|  | World | EU ex-UK exScand Value | EU ex-UK exScand Growth | UK Value | UK Growth | Asia \& Pacific Value | Asia \& Pacific Growth | Scandinavia Value | Scandinavia Growth | US Value | US Growth |
| 1. Mean returns |  |  |  |  |  |  |  |  |  |  |  |
| Regime 1 (Bull Word/Low Vol.) | 0.956** | 1.851** | 1.446** | 1.512** | 1.296** | 1.168** | 0.612* | 1.690** | 1.392** | 1.317** | 1.557** |
| Regime 2 (Bear Word/High Vol.) | 0.049 | 0.036 | -0.209 | 2.247* | 1.494 | 1.623* | -0.118 | 1.643* | 1.716 | 0.501 | 1.182 |
| 2. Correlations/Volatilities |  |  |  |  |  |  |  |  |  |  |  |
| Regime 1 (Bull Word/Low Vol.) |  |  |  |  |  |  |  |  |  |  |  |
| World | 2.910** |  |  |  |  |  |  |  |  |  |  |
| EU ex-UK ex-Scand Value | 0.714** | 4.128** |  |  |  |  |  |  |  |  |  |
| EU ex-UK ex-Scand Growth | 0.787** | 0.860** | 3.268** |  |  |  |  |  |  |  |  |
| United Kingdom Value | 0.613** | 0.562** | 0.557** | 4.431** |  |  |  |  |  |  |  |
| United Kingdom Growth | 0.636** | 0.490** | 0.552** | 0.760** | 3.961** |  |  |  |  |  |  |
| Asia \& Pacific Value | 0.605** | 0.432** | 0.466** | 0.335* | 0.306* | 3.886** |  |  |  |  |  |
| Asia \& Pacific Growth | 0.720** | 0.425** | 0.501** | 0.439** | 0.410** | 0.718** | 3.997** |  |  |  |  |
| Scandinavia Value | 0.507** | 0.532** | 0.466** | 0.496** | 0.431** | 0.383* | 0.394** | 5.387** |  |  |  |
| Scandinavia Growth | 0.595** | 0.495** | 0.518** | 0.446** | 0.457* | 0.282* | 0.385** | 0.651** | 4.740** |  |  |
| United States Value | 0.844** | 0.454** | 0.574** | 0.389** | 0.493** | 0.325** | 0.459** | 0.319* | 0.510** | 3.664** |  |
| United States Growth | 0.817** | 0.614** | 0.626** | 0.518** | 0.517** | 0.327** | 0.420** | 0.403** | 0.507** | 0.786** | 3.133** |
| Regime 2 (Bear Word/High Vol.) |  |  |  |  |  |  |  |  |  |  |  |
| World | 5.438** |  |  |  |  |  |  |  |  |  |  |
| EU ex-UK ex-Scand Value | 0.777** | 5.767** |  |  |  |  |  |  |  |  |  |
| EU ex-UK ex-Scand Growth | 0.791** | 0.845** | 5.948** |  |  |  |  |  |  |  |  |
| United Kingdom Value | 0.663** | 0.637** | 0.576** | 9.131** |  |  |  |  |  |  |  |
| United Kingdom Growth | 0.661** | 0.563** | 0.628** | 0.801** | 8.915** |  |  |  |  |  |  |
| Asia \& Pacific Value | 0.598** | 0.544** | 0.435** | 0.446** | 0.324* | 7.297** |  |  |  |  |  |
| Asia \& Pacific Growth | 0.648** | 0.454** | 0.518** | 0.365* | 0.331* | 0.605** | 6.663** |  |  |  |  |
| Scandinavia Value | 0.550** | 0.663** | 0.543** | 0.369** | 0.338** | 0.440** | 0.298* | 8.436** |  |  |  |
| Scandinavia Growth | 0.656** | 0.622** | 0.716** | 0.333* | 0.380* | 0.365** | 0.514** | 0.642** | 8.959** |  |  |
| United States Value | 0.868** | 0.609** | 0.651** | 0.524** | 0.587** | 0.353* | 0.363* | 0.422** | 0.577** | 6.705** |  |
| United States Growth | 0.797** | 0.663** | 0.577** | 0.584** | 0.591** | 0.386* | 0.299* | 0.460** | 0.389** | 0.785** | 6.293** |
| 3. Transition probabilities |  | Regime 1 ( | Bull Word/Low | olatility) |  |  | Regime | 2 (Bear Word/ | High Volatilit |  |  |
| Regime 1 (Bull Word/Low Vol.) |  |  | 0.901** |  |  |  |  | 0.099 |  |  |  |
| Regime 2 (Bear Word/High Volat |  |  | 0.292 |  |  |  |  | 0.708** |  |  |  |
|  | Panel C - MARKOV CHAIN PROPERTIES, TWO-STATE MODEL |  |  |  |  |  |  |  |  |  |  |
|  | Regime 1 |  | Regime 2 |  | Average duration (in months) |  |  |  |  | Regime 1 | Regime 2 |
| Ergodic Probs | 0.747 |  | 0.253 |  |  |  |  |  |  | 10.10 | 3.43 |

## Table 4

## Moments Implied by Estimated Two-State Markov Switching Model

This table reports moment implied by the estimated Two-State Model for returns. In the co-skewness matrix, coefficients above the main diagonal refer to the sample covariance between the square of the returns of the row portfolio/index and the level of returns of the column portfolio/index; coefficients below the main diagonal refer to the sample covariance between the level of the returns of the column portfolio/index and the square of returns of the row portfolio/index. In the correlation/co-kurtosis matrix, correlations are reported above the main diagonal and sample covariances between squared portfolio returns appear below the main diagonal. Panels A, B and C respectively refer to the International, the Industry and the International Book-to-Market Portfolios.
Panel A (International MSCI USD Returns, 1988:01-2008:08)

|  | Mean | St. Dev. | Sharpe ratio | Median | Min. | Max. | Skewness | Kurtosis |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Pacific ex-Japan | 0.713 | 4.381 | 0.082 | 0.830 | -21.8 | 17.8 | -0.308 | 4.222 |
| Japan | -0.060 | 5.084 | -0.081 | 0.037 | -12.1 | 12.3 | -0.058 | 3.592 |
| Europe ex-UK | 0.748 | 3.964 | 0.100 | 0.818 | -13.1 | 12.2 | -0.247 | 3.672 |
| United Kingdom | 0.703 | 3.558 | 0.098 | 0.723 | -9.8 | 13.4 | -0.005 | 3.279 |
| North America | 0.719 | 3.192 | 0.115 | 0.708 | -9.5 | 9.1 | -0.135 | 3.442 |
| EM Latin America | 1.889 | 7.237 | 0.212 | 2.066 | -27.3 | 21.9 | -0.286 | 4.106 |
| EM Asia | 0.708 | 5.747 | 0.062 | 0.771 | -19.7 | 17.0 | -0.100 | 3.789 |
| EM Europe \& Middle East | 1.843 | 6.290 | 0.237 | 1.875 | -22.1 | 22.0 | -0.055 | 4.101 |


|  | Correlation and (variance) co-kurtosis matrices |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Pacific ex- $\mathrm{JP}$ | Japan | EU ex-UK | UK | North <br> Amr. | EM Latin Amr. | EM Asia | EM EU \& Middle <br> East |
| Pacific ex-JP |  | 0.447 | 0.594 | 0.621 | 0.603 | 0.545 | 0.784 | 0.426 |
| Japan | 1.457 |  | 0.459 | 0.476 | 0.369 | 0.318 | 0.416 | 0.223 |
| EU ex-UK | 1.873 | 1.635 |  | 0.740 | 0.671 | 0.417 | 0.513 | 0.469 |
| UK | 1.731 | 1.494 | 2.241 |  | 0.666 | 0.397 | 0.445 | 0.353 |
| North Amr. | 1.826 | 1.313 | 2.333 | 2.001 |  | 0.501 | 0.552 | 0.401 |
| EM Latin Amr. | 1.862 | 1.350 | 1.788 | 1.322 | 1.861 |  | 0.492 | 0.480 |
| EM Asia | 2.833 | 1.671 | 1.960 | 1.430 | 1.799 | 1.716 |  | 0.481 |
| EM EU \& Middle East | 1.428 | 1.253 | 1.665 | 1.272 | 1.732 | 1.914 | 1.535 |  |
|  | Co-skewness matrix |  |  |  |  |  |  |  |
|  | Pacific exJP | Japan | EU ex-UK | UK | North Amr. | EM Latin Amr. | EM Asia | EM EU \& Middle <br> East |
| Pacific ex-JP |  | -0.164 | -0.195 | -0.124 | -0.178 | -0.210 | -0.213 | -0.166 |
| Japan | -0.104 |  | -0.062 | -0.008 | -0.054 | -0.152 | -0.084 | -0.048 |
| EU ex-UK | -0.193 | -0.168 |  | -0.181 | -0.227 | -0.164 | -0.212 | -0.172 |
| UK | -0.030 | -0.055 | -0.092 |  | -0.021 | -0.055 | -0.048 | -0.090 |
| North Amr. | -0.172 | -0.107 | -0.219 | -0.093 |  | -0.203 | -0.175 | -0.186 |
| EM Latin Amr. | -0.218 | -0.215 | -0.112 | -0.144 | -0.180 |  | -0.135 | -0.179 |
| EM Asia | -0.170 | -0.142 | -0.182 | -0.111 | -0.136 | -0.142 |  | -0.154 |
| EM EU \& Middle East | -0.103 | -0.045 | -0.068 | -0.114 | -0.163 | -0.247 | -0.114 |  |

Table 4 (cont'ed)
Moments Implied by Estimated Two-State Markov Switching Model

Panel B (CRSP Industry Returns, 1926:07-2008:07)

|  |  | Mean | Std. Dev. | Sharpe ratio | Median | Min. | Max. | Skewness | Kurtosis |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Non-Durables |  | 0.969 | 4.198 | 0.158 | 1.050 | -20.7 | 21.0 | -0.147 | 6.187 |
| Durables |  | 1.056 | 6.759 | 0.111 | 1.104 | -31.2 | 40.6 | 0.338 | 10.426 |
| Manufacture |  | 1.017 | 5.631 | 0.126 | 1.074 | -24.8 | 33.3 | 0.263 | 9.314 |
| Energy |  | 1.095 | 5.369 | 0.147 | 1.147 | -18.5 | 27.5 | 0.034 | 5.315 |
| Hi Tech |  | 1.086 | 6.624 | 0.118 | 1.139 | -27.0 | 37.3 | 0.036 | 7.112 |
| Telecom |  | 0.827 | 4.114 | 0.127 | 0.908 | -11.0 | 18.4 | -0.147 | 5.761 |
| Shops/ Distrib. |  | 0.965 | 5.243 | 0.126 | 1.047 | -25.8 | 26.0 | -0.126 | 6.472 |
| Health |  | 1.069 | 5.149 | 0.148 | 1.138 | -23.8 | 24.0 | -0.037 | 6.688 |
| Utilities |  | 0.891 | 5.113 | 0.115 | 0.959 | -15.8 | 24.6 | -0.093 | 8.157 |
| Other |  | 0.905 | 5.792 | 0.104 | 1.003 | -19.6 | 31.5 | 0.174 | 9.673 |
| Correlation and (variance) co-kurtosis matrices |  |  |  |  |  |  |  |  |  |
|  | NonDurables | Durables | Manuf. Energy | Hi Tech | Telecom | Shops/ Distrib. | Health | Utilities | Other |
| Non-Durables |  | 0.754 | 0.8490 .616 | 0.732 | 0.671 | 0.866 | 0.803 | 0.709 | 0.846 |
| Durables | 6.005 |  | 0.8730 .610 | 0.778 | 0.619 | 0.799 | 0.652 | 0.636 | 0.802 |
| Manufacture | 6.145 | 8.597 | 0.728 | 0.861 | 0.673 | 0.840 | 0.760 | 0.706 | 0.904 |
| Energy | 3.664 | 4.456 | 4.849 | 0.609 | 0.501 | 0.577 | 0.562 | 0.616 | 0.690 |
| Hi Tech | 4.869 | 6.595 | 6.7493 .806 |  | 0.677 | 0.785 | 0.721 | 0.627 | 0.797 |
| Telecom | 3.561 | 4.013 | $4.341 \quad 2.744$ | 4.311 |  | 0.672 | 0.600 | 0.636 | 0.695 |
| Shops/ Distrib. | 5.285 | 6.052 | 6.1303 .530 | 5.211 | 3.849 |  | 0.744 | 0.660 | 0.823 |
| Health | 5.008 | 5.526 | 5.8043 .432 | 4.925 | 3.582 | 4.909 |  | 0.629 | 0.739 |
| Utilities | 4.607 | 5.344 | 5.8313 .696 | 5.077 | 4.230 | 4.999 | 4.822 |  | 0.740 |
| Other | 5.535 | 6.608 | $7.628 \quad 4.405$ | 6.082 | 4.707 | 5.749 | 5.651 | 6.344 |  |
| Co-skewness matrix |  |  |  |  |  |  |  |  |  |
|  | NonDurables | Durables | Manuf. Energy | Hi Tech | Telecom | Shops/ Distrib. | Health | Utilities | Other |
| Non-Durables |  | -0.013 | -0.036 -0.092 | -0.069 | -0.147 | -0.132 | -0.101 | -0.129 | -0.101 |
| Durables | 0.126 |  | $0.294 \quad 0.145$ | 0.199 | -0.025 | 0.115 | 0.143 | 0.063 | 0.144 |
| Manufacture | 0.082 | 0.274 | 0.111 | 0.174 | 0.006 | 0.082 | 0.136 | 0.086 | 0.178 |
| Energy | -0.067 | 0.046 | 0.040 | 0.005 | -0.071 | -0.081 | -0.012 | -0.039 | 0.003 |
| Hi Tech | -0.023 | 0.099 | 0.0990 .015 |  | -0.089 | -0.016 | 0.014 | -0.021 | 0.037 |
| Telecom | -0.143 | -0.129 | -0.087 -0.115 | -0.119 |  | -0.147 | -0.107 | -0.087 | -0.085 |
| Shops/ Distrib. | -0.127 | -0.010 | -0.025 -0.092 | -0.060 | -0.135 |  | -0.103 | -0.104 | -0.067 |
| Health | -0.066 | 0.042 | $0.055-0.010$ | -0.004 | -0.066 | -0.071 |  | -0.064 | 0.021 |
| Utilities | -0.128 | -0.039 | -0.007 -0.078 | -0.053 | -0.079 | -0.114 | -0.091 |  | -0.014 |
| Other | -0.011 | 0.125 | 0.1650 .048 | 0.095 | 0.010 | -0.122 | 0.088 | 0.077 |  |

Table 4 (cont'ed)
Moments Implied by Estimated Two-State Markov Switching Model

|  |  |  | Mean | St. Dev. | Sharpe ratio | Median | Min. | Max. | Skewness | Kurtosis |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | World |  | 0.717 | 3.244 | 0.072 | 0.764 | -18.2 | 14.6 | -0.440 | 5.207 |  |
|  | EU ex-UK ex-Scand Value |  | 1.391 | 4.055 | 0.224 | 1.487 | -11.4 | 15.7 | -0.269 | 4.122 |  |
|  | EU ex-UK ex-Scand Growth |  | 1.024 | 3.619 | 0.149 | 1.131 | -15.2 | 15.9 | -0.441 | 5.205 |  |
|  | United Kingdom Value |  | 1.689 | 5.171 | 0.233 | 1.609 | -19.4 | 36.0 | 0.385 | 6.919 |  |
|  | United Kingdom Growth |  | 1.335 | 4.870 | 0.175 | 1.300 | -22.0 | 42.3 | 0.591 | 10.34 |  |
|  | Asia \& Pacific Value |  | 1.285 | 4.338 | 0.184 | 1.253 | -13.0 | 16.0 | 0.040 | 4.894 |  |
|  | Asia \& Pacific Growth |  | 0.405 | 4.172 | -0.019 | 0.441 | -14.5 | 22.0 | -0.060 | 4.418 |  |
|  | Scandinavia Value |  | 1.680 | 5.505 | 0.217 | 1.653 | -12.5 | 24.6 | 0.065 | 4.042 |  |
|  | Scandinavia Growth |  | 1.443 | 5.315 | 0.180 | 1.415 | -21.9 | 26.6 | 0.048 | 4.641 |  |
|  | United States Value |  | 1.091 | 4.028 | 0.151 | 1.143 | -22.4 | 16.3 | -0.247 | 4.549 |  |
|  | United States Growth |  | 1.447 | 3.606 | 0.267 | 1.456 | -15.7 | 21.4 | -0.086 | 5.689 |  |
|  | Correlation and (variance) co-kurtosis matrices |  |  |  |  |  |  |  |  |  |  |
|  | World | EU ex-UK exScand Value |  | UK Value | UK Growth | Asia Pacific | Asia Pacific |  | Scandinavia | United States | United States |
|  |  |  | Scand Growth |  |  | Value | Growth | Value | Growth | Value | Growth |
| World |  | 0.741 | 0.792 | 0.626 | 0.634 | 0.601 | 0.683 | 0.528 | 0.626 | 0.857 | 0.803 |
| EU ex-UK ex-Scand Value | 3.218 |  | 0.851 | 0.567 | 0.496 | 0.478 | 0.442 | 0.580 | 0.538 | 0.524 | 0.623 |
| EU ex-UK ex-Scand Growth | 4.030 | 3.488 |  | 0.543 | 0.574 | 0.447 | 0.513 | 0.498 | 0.613 | 0.617 | 0.594 |
| United Kingdom Value | 3.324 | 2.427 | 3.011 |  | 0.779 | 0.399 | 0.386 | 0.424 | 0.379 | 0.453 | 0.548 |
| United Kingdom Growth | 3.633 | 2.402 | 3.477 | 6.853 |  | 0.315 | 0.350 | 0.370 | 0.408 | 0.535 | 0.552 |
| Asia \& Pacific Value | 3.021 | 2.237 | 2.677 | 2.357 | 2.170 |  | 0.653 | 0.416 | 0.337 | 0.345 | 0.365 |
| Asia \& Pacific Growth | 2.599 | 1.742 | 2.217 | 1.740 | 1.716 | 2.572 |  | 0.352 | 0.454 | 0.413 | 0.357 |
| Scandinavia Value | 2.245 | 2.137 | 2.204 | 1.833 | 1.822 | 1.902 | 1.521 |  | 0.647 | 0.371 | 0.431 |
| Scandinavia Growth | 2.749 | 2.118 | 2.866 | 1.883 | 2.034 | 1.864 | 1.920 | 2.254 |  | 0.544 | 0.442 |
| United States Value | 3.886 | 2.328 | 3.088 | 2.513 | 2.836 | 2.128 | 1.743 | 1.820 | 2.403 |  | 0.783 |
| United States Growth | 3.964 | 2.700 | 3.165 | 3.677 | 4.441 | 2.220 | 1.645 | 1.936 | 1.993 | 4.824 |  |
|  | Co-Skewness matrices |  |  |  |  |  |  |  |  |  |  |
|  | World | EU ex-UK exScand Value | EU ex-UK exScand Growth | UK Value | UK Growth | Asia Pacific Value | Asia Pacific Growth | Scandinavia Value | Scandinavia Growth | United States Value | United States Growth |
| World |  | -0.382 | -0.439 | -0.193 | -0.205 | -0.295 | -0.250 | -0.244 | -0.261 | -0.350 | -0.306 |
| EU ex-UK ex-Scand Value | -0.316 |  | -0.306 | -0.177 | -0.195 | -0.215 | -0.162 | -0.204 | -0.205 | -0.268 | -0.260 |
| EU ex-UK ex-Scand Growth | -0.425 | -0.363 |  | -0.228 | -0.247 | -0.291 | -0.232 | -0.270 | -0.273 | -0.367 | -0.326 |
| United Kingdom Value | -0.018 | -0.084 | -0.078 |  | 0.397 | 0.014 | -0.025 | -0.068 | -0.044 | -0.034 | 0.137 |
| United Kingdom Growth | -0.072 | -0.078 | -0.049 | 0.470 |  | -0.022 | -0.031 | -0.042 | 0.004 | -0.016 | 0.203 |
| Asia \& Pacific Value | -0.204 | -0.179 | -0.222 | -0.044 | -0.109 |  | -0.076 | -0.125 | -0.122 | -0.186 | -0.168 |
| Asia \& Pacific Growth | -0.153 | -0.132 | -0.141 | -0.030 | -0.055 | -0.055 |  | -0.108 | -0.116 | -0.122 | -0.108 |
| Scandinavia Value | -0.135 | -0.147 | -0.160 | -0.049 | -0.082 | -0.059 | -0.050 |  | -0.023 | -0.119 | -0.105 |
| Scandinavia Growth | -0.147 | -0.166 | -0.162 | -0.070 | -0.059 | -0.105 | -0.085 | -0.027 |  | -0.086 | -0.107 |
| United States Value | -0.295 | -0.298 | -0.356 | -0.130 | -0.165 | -0.193 | -0.158 | -0.163 | -0.164 |  | -0.202 |
| United States Growth | -0.216 | -0.266 | -0.298 | 0.025 | 0.004 | -0.150 | -0.139 | -0.151 | -0.134 | -0.179 |  |

Table 5
Portfolio Weights as a Function of the Initial State
This table displays average optimal portfolio shares. The out-of-sample period for our International (Panel A), Industry (Panel B) and Book-to-Market International (Panel C) data runs from 1998:01-2008:07, 1980:01-2008:07, and 1995:01-2007:12 respectively. The 1.h.s. (r.h.s) refers to portfolios subject (free) from short-sales constraints. The first six columns refer to the investor horizon. The upper (lower) part of each panel refers to the allocation associated with the single-state (two-state) model. In the latter case, we highlight the "ex-ante" portfolio shares computed using the ergodic probabilities, and the shares conditional on the bear and the bull states. Each row, within a given case, is associated with investor preferences ranging from mean-variance to four-moments.
Panel A (International MSCI USD Returns, 1988:01-2008:08)


Panel A (International MSCI USD Returns, 1988:01-2008:08)


# Table 5 (cont'ed) 

Portfolio Weights as a Function of the Initial State
Panel B (CRSP Industry Returns, 1926:07-2008:07)

|  |  | T=1 | T=3 | T=12 | T=24 | T=60 | T=120 | "Slope" | T=1 | T=3 | $\mathrm{T}=12$ | T=24 | $\mathrm{T}=60$ | $\mathrm{T}=120$ | "Slope" |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | No-short sales |  |  |  |  |  |  | Unconstrained |  |  |  |  |  |  |
|  |  | Single-State Model (Unconditional Allocation) |  |  |  |  |  |  |  |  |  |  |  |  |  |
| $\backslash$ | Non Durables | 0.106 | 0.000 | 0.000 | 0.095 | 0.159 | 0.194 | 0.088 | 0.868 | 1.055 | 0.820 | 0.846 | 0.834 | 0.838 | -0.031 |
| , | Durables | 0.000 | 0.000 | 0.000 | 0.021 | 0.000 | 0.000 | 0.000 | 0.288 | 0.446 | 0.304 | 0.145 | 0.039 | 0.046 | -0.242 |
| , | Manufacturing | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.154 | -0.104 | -0.138 | -0.134 | -0.129 | -0.138 | 0.016 |
|  | Energy | 0.461 | 0.503 | 0.585 | 0.653 | 0.596 | 0.604 | 0.143 | 0.699 | 1.073 | 0.506 | 0.404 | 0.407 | 0.411 | -0.288 |
|  | Hi Tech | 0.000 | 0.000 | 0.145 | 0.070 | 0.045 | 0.040 | 0.040 | 0.213 | 0.459 | 0.259 | 0.022 | 0.029 | 0.025 | -0.188 |
| $\rangle$ | Telecommunications | 0.000 | 0.000 | 0.000 | 0.000 | 0.115 | 0.113 | 0.113 | -0.090 | -0.632 | -0.140 | 0.160 | 0.255 | 0.309 | 0.399 |
|  | Shops/Distribution | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.172 | -0.351 | -0.107 | -0.110 | -0.103 | -0.099 | 0.073 |
|  | Health | 0.433 | 0.497 | 0.177 | 0.161 | 0.085 | 0.049 | -0.384 | 0.544 | 0.918 | 0.326 | 0.274 | 0.284 | 0.227 | -0.317 |
|  | Utilities | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.174 | -0.266 | -0.080 | -0.013 | -0.014 | -0.014 | 0.160 |
|  | Other | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.023 | -1.598 | -0.749 | -0.593 | -0.602 | -0.606 | 0.417 |
|  |  | Two-State Model (Current State: Ergodic/Unconditional Probabilities) |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  | 0.034 | 0.143 | 0.359 | 0.564 | 0.673 | 0.678 | 0.644 | 0.819 | 1.104 | 0.851 | 0.875 | 0.857 | 0.880 | 0.061 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.195 | 0.468 | 0.095 | 0.104 | 0.106 | 0.106 | -0.089 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.139 | -0.947 | -0.474 | -0.516 | -0.510 | -0.526 | -0.387 |
|  |  | 0.463 | 0.393 | 0.314 | 0.290 | 0.261 | 0.261 | -0.202 | 0.718 | 1.095 | 0.448 | 0.446 | 0.443 | 0.446 | -0.271 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.189 | 0.459 | 0.107 | 0.116 | 0.112 | 0.124 | -0.065 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.021 | -0.857 | -0.004 | 0.133 | 0.127 | 0.118 | 0.140 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.155 | 0.059 | 0.117 | 0.113 | 0.128 | 0.128 | 0.284 |
|  |  | 0.503 | 0.464 | 0.327 | 0.146 | 0.066 | 0.061 | -0.442 | 0.612 | 1.076 | 0.398 | 0.251 | 0.249 | 0.245 | -0.368 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.167 | -0.462 | -0.047 | -0.050 | -0.037 | -0.050 | 0.116 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.050 | -0.995 | -0.491 | -0.471 | -0.474 | -0.471 | 0.579 |
|  | Non Durables | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.326 | 0.616 | 0.804 | 1.016 | 0.977 | 0.986 | 0.660 |
|  | Durables | 0.046 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.046 | 0.319 | 0.440 | 0.109 | 0.106 | 0.110 | 0.103 | -0.216 |
|  | Manufacturing | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.252 | -0.753 | -0.538 | -0.560 | -0.581 | -0.621 | -0.369 |
|  | Energy | 0.954 | 0.709 | 0.581 | 0.345 | 0.314 | 0.305 | -0.649 | 0.670 | 0.816 | 0.483 | 0.407 | 0.445 | 0.536 | -0.134 |
|  | Hi Tech | 0.000 | 0.038 | 0.145 | 0.290 | 0.345 | 0.367 | 0.367 | 0.117 | 0.307 | 0.126 | 0.105 | 0.091 | 0.106 | -0.011 |
|  | Telecommunications | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.181 | -0.381 | 0.083 | 0.151 | 0.131 | 0.092 | -0.089 |
|  | Shops/Distribution | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.076 | 0.022 | 0.092 | 0.117 | 0.116 | 0.143 | 0.219 |
|  | Health | 0.000 | 0.253 | 0.274 | 0.365 | 0.341 | 0.328 | 0.328 | 0.328 | 0.626 | 0.486 | 0.287 | 0.225 | 0.173 | -0.155 |
|  | Utilities | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.022 | -0.123 | -0.016 | -0.043 | -0.052 | -0.045 | -0.023 |
|  | Other | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.591 | -0.569 | -0.630 | -0.587 | -0.461 | -0.473 | 0.118 |
| $\begin{gathered} \text { Mean-Var-Kurtosis } \\ \text { Preferences } \end{gathered}$ | Non Durables | 0.054 | 0.000 | 0.145 | 0.356 | 0.679 | 0.686 | 0.632 | 0.255 | 0.456 | 0.879 | 0.884 | 0.867 | 0.871 | 0.616 |
|  | Durables | 0.000 | 0.053 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.264 | 0.224 | 0.104 | 0.095 | 0.110 | 0.107 | -0.157 |
|  | Manufacturing | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.619 | -0.420 | -0.475 | -0.504 | -0.524 | -0.519 | 0.100 |
|  | Energy | 0.358 | 0.358 | 0.335 | 0.300 | 0.259 | 0.260 | -0.097 | 0.470 | 0.461 | 0.440 | 0.446 | 0.445 | 0.441 | -0.029 |
|  | Hi Tech | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.089 | 0.165 | 0.110 | 0.120 | 0.122 | 0.119 | 0.030 |
|  | Telecommunications | 0.297 | 0.210 | 0.175 | 0.098 | 0.000 | 0.000 | -0.297 | 0.294 | 0.201 | 0.138 | 0.139 | 0.125 | 0.125 | -0.169 |
|  | Shops/Distribution | $0.000$ | 0.047 | 0.040 | 0.024 | 0.000 | 0.000 | $0.000$ | 0.029 | 0.030 | 0.097 | 0.108 | 0.128 | 0.132 | 0.102 |
|  | Health | 0.272 | 0.323 | 0.305 | 0.222 | 0.062 | 0.054 | -0.218 | 0.272 | 0.171 | 0.227 | 0.235 | 0.248 | 0.244 | -0.027 |
|  | Utilities | 0.019 | 0.009 | 0.000 | 0.000 | 0.000 | 0.000 | -0.019 | 0.057 | 0.014 | -0.031 | -0.051 | -0.044 | -0.043 | -0.101 |
|  | Other | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.111 | -0.300 | -0.489 | -0.472 | -0.477 | -0.477 | -0.366 |
|  | Non Durables | 0.000 | 0.000 | 0.144 | 0.405 | 0.663 | 0.679 | 0.679 | 0.094 | 0.156 | 0.460 | 0.886 | 0.873 | 0.859 | 0.766 |
|  | Durables | 0.046 | 0.079 | 0.032 | 0.000 | 0.000 | 0.000 | -0.046 | 0.327 | 0.316 | 0.185 | 0.093 | 0.101 | 0.108 | -0.219 |
|  | Manufacturing | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.457 | -0.405 | -0.465 | -0.513 | -0.505 | -0.525 | -0.067 |
|  | Energy | 0.362 | 0.414 | 0.360 | 0.297 | 0.254 | 0.258 | -0.104 | 0.477 | 0.531 | 0.485 | 0.452 | 0.450 | 0.451 | -0.027 |
|  | Hi Tech | 0.000 | 0.016 | 0.018 | 0.045 | 0.005 | 0.000 | 0.000 | 0.061 | 0.149 | 0.121 | 0.114 | 0.114 | 0.112 | 0.051 |
|  | Telecommunications | 0.286 | 0.151 | 0.084 | 0.045 | 0.000 | 0.000 | -0.286 | 0.326 | 0.184 | 0.159 | 0.134 | 0.126 | 0.127 | -0.199 |
|  | Shops/Distribution | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.047 | -0.037 | 0.078 | 0.129 | 0.131 | 0.138 | 0.185 |
|  | Health | 0.291 | 0.339 | 0.362 | 0.208 | 0.078 | 0.063 | -0.228 | 0.343 | 0.379 | 0.463 | 0.231 | 0.233 | 0.250 | -0.093 |
|  | Utilities | 0.015 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.015 | 0.069 | 0.078 | 0.004 | -0.049 | -0.051 | -0.051 | -0.120 |
|  | Other | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.193 | -0.350 | -0.490 | -0.477 | -0.471 | -0.470 | -0.276 |
|  |  | Two-State Model (Current State: Bear/High Volatility) |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  | 0.000 | 0.000 | 0.054 | 0.267 | 0.405 | 0.516 | 0.516 | -0.045 | -0.217 | 0.632 | 0.748 | 0.844 | 0.945 | 0.990 |
|  |  | 0.000 | 0.000 | 0.056 | 0.103 | 0.094 | 0.000 | 0.000 | 0.188 | 0.199 | 0.002 | 0.004 | 0.003 | 0.007 | -0.181 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.023 | 0.026 | 0.026 | 1.662 | 2.059 | 0.278 | 0.150 | 0.009 | -0.140 | -1.802 |
|  |  | 0.612 | 0.817 | 0.679 | 0.397 | 0.251 | 0.253 | -0.359 | 0.713 | 0.958 | 0.506 | 0.380 | 0.378 | 0.380 | -0.333 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.329 | -0.048 | -0.130 | -0.126 | -0.117 | -0.122 | 0.207 |
|  |  | 0.239 | 0.000 | 0.000 | 0.000 | 0.000 | 0.005 | -0.234 | 0.852 | 0.413 | 0.549 | 0.478 | 0.406 | 0.367 | -0.485 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.144 | -0.730 | -0.358 | -0.354 | -0.306 | -0.253 | 0.891 |
|  |  | 0.148 | 0.183 | 0.211 | 0.233 | 0.227 | 0.201 | 0.053 | 0.251 | 0.109 | 0.209 | 0.310 | 0.330 | 0.341 | 0.090 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.246 | 0.194 | 0.065 | 0.060 | 0.054 | 0.024 | -0.222 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.393 | -1.936 | -0.752 | -0.650 | -0.600 | -0.549 | 0.844 |
|  | Non DurablesDurablesManufacturingEnergyHi TechTelecommunicationShops/DistributionHealthUtilitiesOther | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.010 | -0.056 | 0.401 | 0.576 | 1.079 | 1.314 | 1.304 |
|  |  | 0.246 | 0.088 | 0.113 | 0.224 | 0.184 | 0.145 | -0.101 | 0.290 | 0.328 | 0.103 | 0.062 | 0.007 | -0.024 | -0.314 |
|  |  | 0.000 | 0.025 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.962 | 1.779 | 0.298 | 0.265 | 0.185 | 0.015 | -0.947 |
|  |  | 0.209 | 0.766 | 0.804 | 0.676 | 0.603 | 0.559 | 0.350 | 0.550 | 0.943 | 0.427 | 0.451 | 0.470 | 0.297 | -0.253 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.210 | -0.014 | -0.100 | -0.139 | -0.135 | -0.155 | 0.055 |
|  |  | 0.000 | 0.000 | 0.039 | 0.000 | 0.000 | 0.000 | 0.000 | 0.855 | 0.482 | 0.608 | 0.570 | 0.485 | 0.559 | -0.296 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.931 | -0.779 | -0.477 | -0.287 | -0.296 | -0.214 | 0.717 |
|  |  | 0.545 | 0.121 | 0.001 | 0.066 | 0.203 | 0.296 | -0.249 | 0.213 | 0.010 | 0.057 | 0.242 | 0.172 | 0.272 | 0.059 |
|  |  | 0.000 | 0.000 | 0.043 | 0.034 | 0.010 | 0.000 | 0.000 | 0.216 | 0.196 | 0.084 | 0.075 | 0.073 | 0.004 | -0.212 |
|  |  | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.038 | -1.621 | -0.566 | -0.558 | -0.647 | -0.541 | 0.496 |

Panel B (CRSP Industry Returns, 1926:07-2008:07)


# Table 5 (cont'ed) 

Portfolio Weights as a Function of the Initial State
Panel C (International Book-to-Market Sorted Portfolio Local Returns, 1975:01-2007:12)

|  |  | T=1 | T=3 | T=12 | T=24 | T=60 | T=120 | "Slope" | T=1 | T=3 | T=12 | T=24 | T=60 | T=120 | "Slope" |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | No-short sales |  |  |  |  |  |  | Unconstrained |  |  |  |  |  |  |
|  |  | Single-State Model (Unconditional Allocation) |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | World | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -4.900 | -3.654 | -3.659 | -3.483 | -2.956 | -2.669 | 2.231 |
|  | EU ex-UK ex-Scand Value | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 1.372 | 1.254 | 0.649 | 0.570 | 0.528 | 0.485 | -0.887 |
|  | EU ex-UK ex-Scand Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.910 | -0.954 | 0.018 | 0.495 | 0.869 | 0.984 | 2.894 |
|  | United Kingdom Value | 0.599 | 0.677 | 0.781 | 0.596 | 0.240 | 0.000 | -0.599 | 1.725 | 1.405 | 0.749 | 0.675 | 0.304 | 0.251 | -1.475 |
|  | United Kingdom Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.282 | -0.957 | -0.382 | -0.362 | -0.383 | -0.370 | -0.087 |
|  | Asia \& Pacific Value | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 2.731 | 3.236 | 1.059 | 0.845 | 0.784 | 0.753 | -1.978 |
|  | Asia \& Pacific Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -2.678 | -1.495 | -0.658 | -0.345 | -0.084 | 0.149 | 2.827 |
|  | Scandinavia Value | 0.401 | 0.323 | 0.219 | 0.094 | 0.000 | 0.000 | -0.401 | -0.044 | -0.694 | -0.442 | -0.433 | -0.374 | -0.305 | -0.261 |
|  | Scandinavia Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 2.057 | 1.743 | 1.345 | 0.850 | 0.274 | 0.129 | -1.928 |
|  | United States Value | 0.000 | 0.000 | 0.000 | 0.304 | 0.594 | 0.653 | 0.653 | 0.066 | -1.234 | -0.244 | -0.094 | 0.384 | 0.576 | 0.510 |
|  | United States Growth | 0.000 | 0.000 | 0.000 | 0.006 | 0.166 | 0.347 | 0.347 | 2.862 | 2.351 | 2.565 | 2.282 | 1.654 | 1.017 | -1.846 |
|  |  |  |  |  | wo-s | M | Cu | Sta | dic | ond | nal P | bilit |  |  |  |
|  | World | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -5.059 | -3.954 | -3.145 | -2.759 | -2.460 | -2.049 | 3.010 |
|  | EU ex-UK ex-Scand Value | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 1.347 | 1.848 | 0.897 | 0.845 | 0.794 | 0.746 | -0.601 |
|  | EU ex-UK ex-Scand Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.934 | -1.450 | -0.748 | -0.240 | 0.004 | 0.137 | 2.070 |
| ¢ | United Kingdom Value | 0.649 | 0.633 | 0.329 | 0.085 | 0.000 | 0.000 | -0.649 | 1.728 | 1.143 | 0.648 | 0.539 | 0.375 | 0.204 | -1.524 |
| $\text { . } \frac{\bar{T}}{\omega}$ | United Kingdom Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.290 | -1.238 | -0.847 | -0.704 | -0.633 | -0.516 | -0.226 |
| $\frac{\pi}{10}$ | Asia \& Pacific Value | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 2.754 | 2.095 | 1.560 | 1.048 | 0.986 | 0.934 | -1.820 |
| $\frac{1}{\pi}$ | Asia \& Pacific Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -3.113 | -1.795 | -1.059 | -0.745 | -0.548 | -0.375 | 2.738 |
|  | Scandinavia Value | 0.351 | 0.367 | 0.671 | 0.405 | 0.047 | 0.000 | -0.351 | -0.074 | -0.495 | -0.565 | -0.575 | -0.598 | -0.603 | -0.529 |
|  | Scandinavia Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 1.974 | 1.243 | 0.746 | 0.249 | -0.249 | -0.262 | -2.236 |
|  | United States Value | 0.000 | 0.000 | 0.000 | 0.435 | 0.847 | 0.906 | 0.906 | 0.127 | -0.204 | -0.495 | 0.345 | 0.859 | 1.743 | 1.616 |
|  | United States Growth | 0.000 | 0.000 | 0.000 | 0.075 | 0.106 | 0.094 | 0.094 | 3.540 | 3.807 | 4.008 | 2.997 | 2.469 | 1.041 | -2.499 |
|  | World | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -3.243 | -3.038 | -3.043 | -2.974 | -2.756 | -2.635 | 0.608 |
|  | EU ex-UK ex-Scand Value | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.575 | 0.461 | 0.742 | 0.771 | 0.815 | 0.868 | 0.293 |
|  | EU ex-UK ex-Scand Growt | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.123 | -0.935 | -0.483 | -0.043 | 0.065 | 0.141 | 1.264 |
|  | United Kingdom Value | 1.000 | 1.000 | 0.435 | 0.349 | 0.174 | 0.058 | -0.942 | 1.269 | 1.055 | 0.836 | 0.650 | 0.379 | 0.220 | -1.048 |
| Wै © | United Kingdom Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.055 | -0.608 | -0.536 | -0.454 | -0.339 | -0.218 | -0.163 |
| $\frac{1}{10} \frac{\overline{0}}{0}$ | Asia \& Pacific Value | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 1.899 | 2.293 | 1.304 | 1.145 | 0.951 | 0.886 | -1.013 |
| $\frac{1}{\frac{1}{0}}$ | Asia \& Pacific Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.587 | -1.077 | -0.495 | -0.363 | -0.148 | -0.106 | 1.481 |
| $\stackrel{\pi}{5}$ | Scandinavia Value | 0.000 | 0.000 | 0.259 | 0.506 | 0.648 | 0.704 | 0.704 | -0.055 | -0.288 | -0.458 | -0.532 | -0.600 | -0.634 | -0.578 |
|  | Scandinavia Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 1.177 | 0.623 | 0.346 | 0.050 | -0.242 | -0.295 | -1.472 |
|  | United States Value | 0.000 | 0.000 | 0.000 | 0.059 | 0.150 | 0.230 | 0.230 | 0.465 | 0.503 | 0.646 | 0.982 | 1.918 | 2.085 | 1.620 |
|  | United States Growth | 0.000 | 0.000 | 0.306 | 0.086 | 0.028 | 0.008 | 0.008 | 1.679 | 2.011 | 2.142 | 1.768 | 0.956 | 0.687 | -0.992 |
|  | World | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.745 | -2.469 | -3.059 | -3.596 | -3.857 | -3.950 | -2.205 |
|  | EU ex-UK ex-Scand Value | 0.003 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.003 | 0.246 | 0.720 | 0.881 | 0.875 | 0.890 | 0.897 | 0.651 |
|  | EU ex-UK ex-Scand Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.093 | 0.104 | 0.136 | 0.154 | 0.184 | 0.193 | 0.100 |
| y | United Kingdom Value | 0.097 | 0.343 | 0.294 | 0.154 | 0.035 | 0.000 | -0.097 | 0.171 | 0.676 | 0.150 | 0.067 | 0.049 | 0.028 | -0.143 |
| $\frac{\stackrel{y}{3}}{\substack{0}}$ | United Kingdom Growth | 0.095 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.095 | 0.110 | -0.078 | -0.385 | -0.490 | -0.514 | -0.530 | -0.640 |
| $\frac{\overline{1}}{\frac{1}{0}} \frac{\bar{y}}{4}$ | Asia \& Pacific Value | 0.146 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.146 | 0.244 | 0.705 | 1.015 | 0.999 | 0.989 | 0.998 | 0.754 |
| \% | Asia \& Pacific Growth | 0.169 | 0.084 | 0.000 | 0.000 | 0.000 | 0.000 | -0.169 | 0.492 | 0.227 | 0.214 | 0.209 | 0.206 | 0.195 | -0.297 |
| $\frac{1}{\pi}$ | Scandinavia Value | 0.240 | 0.420 | 0.553 | 0.394 | 0.145 | 0.095 | -0.145 | 0.166 | 0.085 | -0.405 | -0.591 | -0.609 | -0.634 | -0.800 |
| $\mathcal{L}$ | Scandinavia Growth | 0.080 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.080 | 0.102 | 0.035 | -0.182 | -0.209 | -0.208 | -0.242 | -0.345 |
|  | United States Value | 0.078 | 0.000 | 0.000 | 0.384 | 0.794 | 0.884 | 0.806 | 0.782 | 1.084 | 1.843 | 2.045 | 2.305 | 2.475 | 1.693 |
|  | United States Growth | 0.090 | 0.153 | 0.153 | 0.068 | 0.026 | 0.021 | -0.069 | 0.338 | -0.089 | 0.793 | 1.537 | 1.566 | 1.570 | 1.231 |
|  | World | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -2.218 | -3.045 | -3.345 | -3.875 | -3.985 | -4.140 | -1.922 |
| $\cdots$ | EU ex-UK ex-Scand Value | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.045 | 0.212 | 0.670 | 0.748 | 0.894 | 0.913 | 0.958 |
|  | EU ex-UK ex-Scand Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.164 | -0.809 | -0.059 | 0.095 | 0.116 | 0.161 | 0.325 |
| $\frac{\overrightarrow{3}}{1}$ | United Kingdom Value | 0.229 | 0.541 | 0.274 | 0.134 | 0.034 | 0.000 | -0.229 | 0.548 | 1.359 | 0.854 | 0.745 | 0.459 | 0.274 | -0.274 |
| $\sum_{0}^{1}$ | United Kingdom Growth | 0.082 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.082 | 0.197 | 0.082 | -0.145 | -0.257 | -0.450 | -0.491 | -0.688 |
| vi | Asia \& Pacific Value | 0.275 | 0.084 | 0.000 | 0.000 | 0.000 | 0.000 | -0.275 | 0.772 | 1.827 | 1.018 | 1.094 | 1.064 | 1.007 | 0.234 |
| $\frac{\sqrt[1]{5}}{\sqrt{0}}$ | Asia \& Pacific Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.147 | -0.330 | -0.084 | 0.074 | 0.156 | 0.187 | 0.040 |
| ${ }_{1}^{10}$ | Scandinavia Value | 0.151 | 0.259 | 0.515 | 0.294 | 0.114 | 0.037 | -0.114 | -0.041 | -0.175 | -0.305 | -0.475 | -0.567 | -0.607 | -0.566 |
| $\frac{5}{0}$ | Scandinavia Growth | 0.228 | 0.047 | 0.000 | 0.000 | 0.000 | 0.000 | -0.228 | 0.463 | 0.174 | -0.183 | -0.236 | -0.217 | -0.255 | -0.718 |
| $\Sigma$ | United States Value | 0.000 | 0.037 | 0.200 | 0.560 | 0.674 | 0.749 | 0.749 | 0.734 | 1.145 | 1.748 | 1.904 | 2.456 | 2.955 | 2.221 |
|  | United States Growth | 0.035 | 0.032 | 0.011 | 0.012 | 0.178 | 0.214 | 0.179 | 0.606 | 0.693 | 0.830 | 0.894 | 0.953 | 0.997 | 0.390 |
|  |  |  |  |  |  | -State | Model | Current | te: Wo | d Bull/L | w Vola | lity) |  |  |  |
|  | World | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -3.405 | -3.004 | -2.748 | -2.244 | -2.009 | -1.780 | 1.625 |
|  | EU ex-UK ex-Scand Value | 0.807 | 0.495 | 0.245 | 0.094 | 0.035 | 0.024 | -0.783 | 4.277 | 3.884 | 2.364 | 1.303 | 0.984 | 0.967 | -3.310 |
|  | EU ex-UK ex-Scand Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -1.868 | -0.847 | -0.394 | -0.034 | 0.064 | 0.146 | 2.014 |
|  | United Kingdom Value | 0.000 | 0.063 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 1.381 | 0.840 | 0.374 | 0.073 | 0.021 | -0.072 | -1.453 |
| $\frac{\cdot \overline{0}}{\pi}$ | United Kingdom Growth | 0.000 | 0.034 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.909 | -1.450 | -0.563 | -0.602 | -0.549 | -0.548 | 0.361 |
| $\frac{1}{0} \frac{\overline{0}}{0}$ | Asia \& Pacific Value | 0.109 | 0.084 | 0.063 | 0.005 | 0.000 | 0.000 | -0.109 | 2.708 | 1.330 | 1.093 | 1.080 | 1.063 | 1.009 | -1.699 |
| $\frac{1}{\pi}$ | Asia \& Pacific Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -4.535 | -2.048 | -0.749 | -0.394 | -0.093 | 0.074 | 4.609 |
| $\frac{0}{2}$ | Scandinavia Value | 0.074 | 0.056 | 0.000 | 0.000 | 0.000 | 0.000 | -0.074 | 0.762 | 2.096 | -0.727 | -0.734 | -0.696 | -0.718 | -1.480 |
|  | Scandinavia Growth | 0.009 | 0.034 | 0.000 | 0.000 | 0.000 | 0.000 | -0.009 | -0.423 | 0.044 | -0.453 | -0.440 | -0.441 | -0.434 | -0.011 |
|  | United States Value | 0.000 | 0.047 | 0.103 | 0.348 | 0.684 | 0.847 | 0.847 | 2.009 | 0.050 | 3.370 | 3.074 | 2.794 | 1.745 | -0.264 |
|  | United States Growth | 0.001 | 0.187 | 0.589 | 0.553 | 0.281 | 0.129 | 0.128 | 1.002 | 0.693 | -0.567 | 0.894 | 0.953 | 0.611 | -0.391 |
|  | World | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -2.902 | -3.018 | -2.703 | -1.919 | -2.252 | -1.932 | 0.970 |
|  | EU ex-UK ex-Scand Value | 0.204 | 0.603 | 0.365 | 0.184 | 0.145 | 0.084 | -0.120 | 2.012 | 2.614 | 1.869 | 1.052 | 0.850 | 0.560 | -1.452 |
|  | EU ex-UK ex-Scand Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.946 | -0.554 | -0.099 | 0.074 | 0.250 | 0.483 | 1.429 |
| $\underset{\sim}{\omega}$ | United Kingdom Value | 0.000 | 0.084 | 0.345 | 0.249 | 0.135 | 0.074 | 0.074 | 0.643 | 0.705 | 0.288 | 0.077 | -0.019 | -0.304 | -0.947 |
|  | United Kingdom Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.371 | -0.643 | -0.383 | -0.569 | -0.249 | -0.086 | 0.285 |
| $\frac{1}{10} \frac{\overline{0}}{4}$ | Asia \& Pacific Value | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 1.772 | 1.287 | 1.094 | 1.086 | 0.952 | 0.958 | -0.814 |
| $\underset{\frac{c}{\pi}}{\substack{4 \\ \hline}}$ | Asia \& Pacific Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -2.188 | -1.083 | -0.583 | -0.202 | -0.001 | 0.099 | 2.287 |
| © | Scandinavia Value | 0.000 | 0.256 | 0.290 | 0.453 | 0.506 | 0.629 | 0.629 | 0.371 | 0.953 | -0.658 | -0.636 | -0.812 | -0.727 | -1.098 |
|  | Scandinavia Growth | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | -0.165 | 0.103 | -0.594 | -0.272 | -0.339 | -0.409 | -0.244 |
|  | United States Value | 0.000 | 0.000 | 0.000 | 0.084 | 0.194 | 0.204 | 0.204 | 1.580 | 0.677 | 2.323 | 2.675 | 2.480 | 1.999 | 0.419 |
|  | United States Growth | 0.796 | 0.000 | 0.000 | 0.030 | 0.020 | 0.009 | -0.787 | 1.193 | 0.693 | 0.446 | 0.894 | 0.953 | 0.358 | -0.834 |

Panel C (International Book-to-Market Sorted Portfolio Local Returns, 1975:01-2007:12)


## Table 6

## Out-of-Sample Performance

This table reports the best, second best and third best model for stock portfolios, in the case of No Short Sales, over four investment horizons: $\mathrm{T}=1,12,60,120$. The last column reports the performance of the equally weighted strategy. We rank models according to the Sharpe Ratio (SR), the Sortino Ratio and the Certainty Equivalent (CEQ). The symbol * indicates that the model has the same ranking when short sales are allowed.

| Panel A (International) | Best | Second Best | Third Best | 1/N |
| :---: | :---: | :---: | :---: | :---: |
| T=1 |  |  |  |  |
| SR | MV(1) * | MVSK * | MVS * | 1/N |
|  | 0.528 | 0.473 | 0.333 | 0.389(third best) |
|  | [-0.077,1.201] | [-0.125,1.127] | [-0.281,0.990] | [-0.225, 1.039] |
| SO RT | MV(1) * | MVSK * | MVS * | 1/N |
|  | 0.777 | 0.662 | 0.460 | 0.561 (third best) |
|  | [-0.117,1.776] | [-0.174, 1.710] | [-0.376, 1.593 ] | [-0.346,1.521] |
| CEQ | MVSK * | MV(1) * | MV(2) * | 1/N |
|  | 3.910 | 0.183 | -1.344 | 3.291 (second best) |
|  | [-8.668,15.999] | [-17.651,18.236] | [-17.202,13.763] | [-8.095,14.243] |
| $\mathrm{T}=12$ |  |  |  |  |
| SR | MV(1) * | MVSK * | MV(2) | 1/N |
|  | 0.653 | 0.546 | 0.485 | 0.467 |
|  | [0.469,0.865] | [0.362,0.755] | [0.308,0.685] | [0.281,0.675] |
| SO RT | MV(1) | MVSK | MV(2) | 1/N |
|  | 1.244 | 0.925 | 0.841 | 0.843 (third best) |
|  | [1.011,1.524] | [0.695,1.221] | [0.539,1.209] | [0.566,1.079] |
| CEQ | MV(2) | MV(1) | MVSK | 1/N |
|  | 4.993 | 4.713 | 4.580 | 2.947 |
|  | [1.251,9.163] | [-0.132,10.569] | [0.107,9.461] | [-1.254,7.627] |
| $\mathrm{T}=60$ |  |  |  |  |
| SR | MVS | MV(2) | MVSK | 1/N |
|  | 0.671 | 0.667 | 0.496 | 0.374 |
|  | [0.557,0.824] | [0.554,0.819] | [0.411,0.608] | [0.291,0.472] |
| SO RT | MV(1) | MVSK | MV(2) | 1/N |
|  | 1.437 | 1.386 | 1.345 | 1.392 (second best) |
|  | [1.193,1.896] | [1.076,1.954] | [1.060,1.976] | [1.021,1.829] |
| CEQ | MVS | MV(2) | MV(1) | 1/N |
|  | 12.3122 | 12.1958 | 9.6215 | 6.4525 |
|  | [9.799,15.350] | [9.757,15.256] | [6.141,15.086] | [4.643,8.786] |
| $\mathrm{T}=120$ |  |  |  |  |
| SR | MV(1) | MVS | MV(2) | 1/N * |
|  | 1.709 | 1.483 | 1.479 | 2.675 (best) |
| SO RT | MV(1) * | MVS | MV(2) * | 1/N * |
|  | 4.463 | 2.977 | 2.965 | 5.268 (best) |
| CEQ | MV(1) * | MVSK | MVS | 1/N |
|  | 49.921 | 20.330 | 19.573 | 18.728 |

Table 6 (cont'ed)
Out-of-Sample Performance

| Panel B (Industry) | Best | Second Best | Third Best | 1/N |
| :---: | :---: | :---: | :---: | :---: |
| T=1 |  |  |  |  |
| SR | MV(2) | MVSK | MV(1) * | 1/N |
|  | 1.138 | 0.984 | 0.874 | 0.754 |
|  | [0.744,1.553] | [0.634,1.347] | [0.490,1.283] | [0.451,1.244] |
| SO RT | MVSK * | MV(2) | MV(1) * | 1/N |
|  | 1.751 | 1.703 | 1.276 | 1.095 |
|  | [1.079,2.563] | [1.032,2.590] | [0.674,2.059] | [0.609,1.958] |
| CEQ | MV(2) | MVSK | MV(1) * | 1/N |
|  | 15.086 | 13.755 | 10.867 | 12.655 (third best) |
|  | [8.813,21.207] | [7.257,20.415] | [5.209,16.413] | [4.573,15.742] |
| T=12 |  |  |  |  |
| SR | MVSK * | $\mathrm{MV}(1)$ * | MV(2) | 1/N |
|  | 0.943 | 0.919 | 0.901 | 0.750 |
|  | [0.822,1.082] | [0.803,1.048] | [0.794,1.016] | [0.724,0.962] |
| SO RT | MV(2) | MVSK * | MV(1) | 1/N |
|  | 1.745 | 1.570 | 1.527 | 1.395 |
|  | [1.550,1.971] | [1.370,1.789] | [1.335,1.791] | [1.219,1.634] |
| CEQ | MV(2) | MVSK * | MV(1) | $1 / \mathrm{N}$ * |
|  | 12.772 | 12.645 | 12.501 | 11.369 |
|  | [11.137,14.454] | [10.932,14.444] | [10.793,14.248] | [9.604,13.227] |
| T=60 |  |  |  |  |
| SR | MV (1) * | MVS | MV(2) | 1/N |
|  | 0.882 | 0.864 | 0.819 | 0.754 |
|  | [0.803,0.983] | [0.788,0.957] | [0.751,0.904] | [0.689,0.834] |
| SO RT | MVSK | MV(2) | MV(1) | 1/N |
|  | 1.718 | 1.705 | 1.628 | 1.314 |
|  | [1.594,1.874] | [1.587,1.868] | [1.493,1.809] | [1.242,1.414] |
| CEQ | MV(2) | MVSK | MV(1) | 1/N |
|  | 18.511 | 18.158 | 17.828 | 15.326 |
|  | [17.294,19.859] | [16.963,19.476] | [16.561,19.244] | [14.072,16.729] |
| T=120 |  |  |  |  |
| SR | MV(1) * | MVS | MV(2) * | 1/N * |
|  | 1.071 | 0.972 | 0.796 | 0.894 (third best) |
|  | [0.989,1.179] | [0.902,1.061] | [0.737,0.871] | [0.824,0.982] |
| SO RT | MV(1) | MVS | MVSK | 1/N |
|  | 2.184 | 1.897 | 1.878 | 1.655 |
|  | [2.009,2.480] | [1.694,2.174] | [1.703,2.094] | [1.482,1.893] |
| CEQ | MV(1) * | MV(2) | MVS | 1/N |
|  | 30.922 | 30.247 | 30.040 | 27.602 |
|  | [29.832,32.113] | [28.499,32.190] | [28.557,31.719] | [25.793,29.680] |

Table 6 (cont'ed)
Out-of-Sample Performance

| Panel C (Book-to-Market) | Best | Second Best | Third Best | 1/N |
| :---: | :---: | :---: | :---: | :---: |
| T=1 |  |  |  |  |
| SR | MVSK * | MV(2) | MVS* | 1/N |
|  | 1.322 | 1.185 | 0.652 | 0.593 |
|  | [0.741,1.973] | [0.618,1.843] | [0.087,1.265] | [0.032,1.215] |
| SO RT | MVSK * | MV(2) | MVS | 1/N |
|  | 1.857 | 1.586 | 0.897 | 0.790 |
|  | [1.005,2.993] | [0.811,2.689] | [0.122,1.891] | [0.043,1.704] |
| CEQ | MVSK * | MV(2) | MVS | 1/N |
|  | 18.643 | 17.701 | 9.858 | 9.126 |
|  | [11.136,25.988] | [9.127,26.252] | [1.098,18.436] | [1.246,16.785] |
| T=12 |  |  |  |  |
| SR | MVS | MV(2) | MVSK * | 1/N |
|  | 0.606 | 0.546 | 0.543 | 0.542 |
|  | [0.415,0.836] | [0.360,0.770] | [0.367,0.750] | [0.352,0.770] |
| SO RT | MVS | MVSK | MV(2) | 1/N |
|  | 0.791 | 0.751 | 0.666 | 0.686 (third best) |
|  | [0.607,1.028] | [0.577,1.024] | [0.500,0.901] | [0.560,0.882] |
| CEQ | MVS | MVSK | MV(2) | 1/N |
|  | 8.586 | 7.432 | 7.328 | 6.926 |
|  | [5.258,12.138] | [4.053,11.215] | [3.944,10.949] | [3.308,10.802] |
| $\mathrm{T}=60$ |  |  |  |  |
| SR | MVS * | MVSK * | MV(2) * | 1/N |
|  | 0.571 | 0.186 | 0.169 | 0.280 (second best) |
|  | [0.492,0.672] | [0.118,0.251] | [0.097,0.236] | [0.212,0.353] |
| SO RT | MVS * | MVSK * | MV(2) * | 1/N |
|  | 1.253 | 0.811 | 0.711 | 0.964 (second best) |
|  | [1.089,1.571] | [0.487,1.025] | [0.403,0.959] | [0.793,1.186] |
| CEQ | MVS * | MVSK * | MV(2) * | 1/N |
|  | 12.868 | 3.603 | 3.378 | 6.469 (second best) |
|  | [11.278,14.762] | [2.512,4.959] | [2.360,4.575] | [5.319,7.870] |
| T=120 |  |  |  |  |
| SR | MVS | MV(2) | MVSK | 1/N |
|  | 3.836 | 1.069 | 1.007 | 3.451 (second best) |
|  | [3.258,4.953] | [0.902,1.396] | [0.848,1.306] | [2.692,5.323] |
| SO RT | MVS * | MV(2) | MVSK | 1/N |
|  | 6.040 | 2.216 | 2.028 | 3.711 (second best) |
|  | [4.992,9.236] | [1.810,3.284] | [1.659,2.994] | [3.040,7.034] |
| CEQ | MVS * | MVSK * | MV(2) | 1/N |
|  | 30.206 | 16.165 | 14.964 | 20.768 (second best) |
|  | [29.524,30.906] | [15.139,17.363] | [14.135,15.881] | [20.223,21.278] |

## Table 7

This table reports the best, second best and third best model for stock portfolios, in the case of No Short Sales, over four investment horizons: $\mathrm{T}=1,12,60,120$. The last column reports the performance of the equally weighted strategy. We rank models according to the Treynor Ratio (TR) and Jensen's Alfa (JA). The symbol * indicates that the model has the same ranking when short sales are allowed.

| Panel A (International) | Best | Second Best | Third Best | 1/N |
| :---: | :---: | :---: | :---: | :---: |
| T=1 |  |  |  |  |
| TR | MV(1) | MVS | MVSK | 1/N |
|  | 0.367 | 0.146 | 0.126 | 0.130 (third best) |
| JA | [-3.607,3.426] | [-1.690,1.867] | [-0.469,1.022] | [-1.026,1.247] |
|  | MV(1) | MVS | MV(2) | 1/N |
|  | 313.260 | 296.129 | 287.387 | 297.290 (second best) |
|  | [198.878,411.150] | [191.352,393.037] | [205.655,365.680] | [223.143,364.240] |
| $\mathrm{T}=12$ |  |  |  |  |
| TR | MV(1) | MV(2) | MVSK* | 1/N |
|  | 0.140 | 0.109 | 0.103 | 0.089 |
|  | [0.103,0.180] | [0.069,0.152] | [0.070,0.136] | [0.056,0.123] |
| JA | MV(2)* | MVS | MVK | 1/N |
|  | 22.530 | 12.255 | -6.755 | -3.540 (third best) |
|  | [7.441,37.445] | [-2.878,27.317] | [-19.506,5.660] | [-15.744,8.068] |
| $\mathrm{T}=60$ |  |  |  |  |
| TR | MVS | MV(2) | MVSK | 1/N |
|  | 0.284 | 0.282 | 0.190 | 0.139 |
|  | [0.223,0.356] | [0.223,0.354] | [0.146,0.237] | [0.098,0.183] |
| JA | MVS* | MV(2)* | MVSK* | 1/N |
|  | 31.946 | 31.735 | 7.768 | -10.883 |
|  | [24.979,38.988] | [25.047,38.570] | [1.563,13.793] | [-16.591,-5.717] |
| $\mathrm{T}=120$ |  |  |  |  |
| TR | MV(1) | MVK | MVSK | 1/N |
|  | -0.405 | -0.604 | -0.609 | -2.952 |
| JA | MV(1)* | MVK | MVSK | 1/N |
|  | 286.784 | 109.252 | 109.246 | 82.477 |

Table 7
Out-of-Sample Performance

| Panel B (Industry) | Best | Second Best | Third Best | 1/N |
| :---: | :---: | :---: | :---: | :---: |
| $\mathrm{T}=1$ |  |  |  |  |
| TR | MVSK | MV(2) | MVS | 1/N |
|  | 0.080 | 0.055 | 0.054 | 0.037 |
|  | [0.054,0.113] | [0.037,0.073] | [0.025,0.087] | [0.021, 0.054 ] |
| JA | MVSK | MVS | MV(1)* | 1/N |
|  | 124.699 | 108.774 | 73.429 | 35.731 |
|  | [83.205,165.834] | [64.180,155.209] | [51.097,97.218] | [24.945,47.499] |
| $\mathrm{T}=12$ |  |  |  |  |
| TR | MVSK | MV(1) | MV(2) | 1/N |
|  | 0.186 | 0.185 | 0.184 | 0.144 |
|  | [0.161,0.213] | [0.161,0.212] | [0.160,0.208] | [0.126,0.162] |
| JA | MV(1) | MVSK | MV(2) | 1/N |
|  | 31.555 | 31.347 | 28.063 | 12.571 |
|  | [25.616,37.712] | [25.191,37.402] | [21.258,34.635] | [8.878,16.430] |
| $\mathrm{T}=60$ |  |  |  |  |
| TR | MV(1)* | MVSK | MV(2)* | 1/N |
|  | 0.632 | 0.571 | 0.557 | 0.456 |
|  | [0.569,0.709] | [0.504,0.657] | [0.506,0.617] | [0.422,0.491] |
| JA | MV(1)* | MV(2) | MVS | 1/N |
|  | 31.672 | 22.395 | 21.431 | 10.315 |
|  | [26.159,37.330] | [16.472,28.219] | [18.344,24.605] | [7.768,12.835] |
| $\mathrm{T}=120$ |  |  |  |  |
| TR | MV(1)* | MVS | MVSK | 1/N |
|  | 1.5786 | 1.110 | 1.080 | 0.994 |
|  | [1.373,1.839] | [1.033,1.197] | [0.884,1.262] | [0.930,1.063] |
| JA | MV(1)* | MVS | MV(2) | 1/N |
|  | 59.871 | 29.693 | 26.274 | 17.707 |
|  | [49.970,69.109] | [24.127,35.387] | [13.506,38.705] | [13.441,22.102] |

Table 7
Out-of-Sample Performance

| Panel C(BM) | Best | Second Best | Third Best | 1/N |
| :---: | :---: | :---: | :---: | :---: |
| $T=1$ |  |  |  |  |
| TR | MVSK | MV(2) | MVS | 1/N |
|  | 0.055 | 0.054 | 0.029 | 0.023 |
|  | [0.032,0.077] | [0.029,0.081] | [0.004,0.054] | [0.001,0.045] |
| J'S A | MVSK | MV(1) | MV(2) | 1/N |
|  | 39.152 | 32.557 | 29.587 | 16.719 |
|  | [18.426,58.609] | [11.988,55.292] | [-10.620,71.605] | [3.111,29.012] |
| $\mathrm{T}=12$ |  |  |  |  |
| TR | MVS* | MV(2) | MVSK* | 1/N |
|  | 0.114 | 0.099 | 0.094 | 0.089 |
|  | [0.081,0.150] | [0.068,0.131] | [0.066,0.120] | [0.061,0.115] |
| JA | MVS | MV(2) | MV(1) | 1/N |
|  | 22.040 | 17.132 | 5.400 | 2.790 |
|  | [14.044,30.922] | [9.911,25.261] | [1.768,9.213] | [0.337,5.086] |
| $\mathrm{T}=60$ |  |  |  |  |
| TR | MVS* | MVSK | MV(2) | 1/N |
|  | 0.339 | 0.098 | 0.087 | 0.140 (second best) |
|  | [0.266,0.419] | [0.056,0.141] | [0.044,0.130] | [0.096,0.186] |
| JA | MVS* | MV(2)* | MV(1) | 1/N |
|  | 36.935 | -4.299 | -5.287 | 8.321(second best) |
|  | [31.744,41.297] | [-7.003,-0.927] | [-6.729,-3.573] | [7.532,9.108] |
| $T=120$ |  |  |  |  |
| TR | MVS* | MV(2) | MVSK | 1/N |
|  | 0.721 | 0.120 | 0.062 | 0.375 (second best) |
|  | [0.493,1.877] | [0.099,0.138] | [0.031,0.092] | [0.329,0.464] |
| JA | MVS* | MV(1) | MV(2) | 1/N |
|  | 69.354 | -17.285 | -33.823 | 32.067 (second best) |
|  | [52.828,96.072] | [-27.667,-10.887] | [-62.339,-17.071] | [26.540,40.516] |

Figure 1

## Smoothed State Probabilities from Two-State Markov Switching Models International Data

The graphs plot the smoothed state probabilities for the two-state switching model. Panels $\mathrm{A}, \mathrm{B}$ and C respectively refer to the International, the Industry and the International Book-to-Market Portfolios.

Panel A (International MSCI USD Returns, 1988:01-2008:08)



Figure 1 (cont'ed)

## Smoothed State Probabilities from Two-State Markov Switching Models Industry Data

Panel B (CRSP Industry Returns, 1926:07-2008:07)



Figure 1 (cont'ed)
Smoothed State Probabilities from Two-State Markov Switching Models -Book-to-Market International Data

Panel C (International Book-to-Market Sorted Portfolio Local Returns, 1975:01-2007:12)



## Figure 2

Dynamics of Portfolio Weights over Time, No Short Sales Admitted \& CRRA = 5
This figure displays how each portfolio share changes as the probability of being in a bear/bull state is updated by the investor. The dotted/solid/ dashed lines respectively identify investor horizons of 1,12 and 120 months. Colors refer to investor preferences over moments of the return distribution. Panels A, B and C respectively refer to the International, the Industry and the International Book-to-Market Portfolios.

Panel A (International MSCI USD Returns, 1988:01-2008:08)








Emerging Markets, Europe and Middle East


Figure 2 (cont'ed)

## Dynamics of Portfolio Weights over Time, No Short Sales Admitted \& CRRA = 5

Panel B (CRSP Industry Returns, 1926:07-2008:07)









Figure 2 (cont'ed)

## Dynamics of Portfolio Weights over Time, No Short Sales Admitted \& CRRA = 5

Panel C (International Book-to-Market Sorted Portfolio Local Returns, 1975:01-2007:12)











## Figure 3

## Dynamics of Portfolio Weights over Time, CRRA = 5 \& Investment Horizon of 12 Months

This figure displays how each portfolio share changes as the probability of being in a bear/bull state is updated by the investor with a 1-year horizon. The red/green/ pink lines respectively identify investor preferences over mean and variance, mean variance and skew, mean variance and kurtosis, when returns follows a two-state switching model. The blue line depicts the dynamics of the mean-variance allocations when returns follow a single-state model. Colors refer to investor preferences over moments of the return distribution. Panels A, B and C respectively refer to the International, the Industry and the International Book-to-Market Portfolios.

Panel A (International MSCI USD Returns, 1988:01-2008:08)






Emerging Markets, Asia



Emerging Markets, Europe and Middle East


Figure 3 (cont'ed)
Dynamics of Portfolio Weights over Time, CRRA = 5 \& Investment Horizon of 12 Months
Panel B (CRSP Industry Returns, 1926:07-2008:07)











Figure 3 (cont'ed)
Dynamics of Portfolio Weights over Time, CRRA = 5 \& Investment Horizon of 12 Months Panel C (International Book-to-Market Sorted Portfolio Local Returns, 1975:01-2007:12)













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[^1]:    ${ }^{1}$ Allowing for a small amount of aversion to ambiguity about returns also leads to an out-of-sample increase in Sharpe ratio. See Garlappi et al., 2007.
    ${ }^{2}$ Exceptions are Fugazza et al (2009) and Diris et al (2008), who however focus on linear forecasting models.

[^2]:    ${ }^{3}$ This confirms previous specification tests performed by Ang and Chen (2009) and Guidolin and Nicodano (2009), who also extend the comparison to non-linear models such as GARCH-M and EGARCH-VAR. Earlier studies uncover non-normal features in retuns of equity portfolios (Longin and Solnik, 2001) as well as found evidence of regimes (e.g., Ang and Bekaert (2002), Turner, Startz and Nelson (1989)).
    ${ }^{4}$ Numerical techniques such as quadrature methods (Ang and Bekaert (2001), Lynch (2001)) or Monte Carlo simulations (Barberis (2000)) may not be very precise when the return distributions are not Gaussian (Keim and Stambaugh (1986), Fama and French (1988) and Pesaran and Timmermann,1995) as is strongly suggested by empirical research. By contrast, Monte Carlo methods rely on discretization of the state space and use grid that are computationally expensive.

[^3]:    ${ }^{5}$ Our partial equilibrium framework treats returns as exogeneous, as in Ang and Bekart (2001), Barberis (2000), Campbell et al. (2003), Das and Uppal (2004), and Kandel and Stambaugh (1996).
    ${ }^{6}$ Classic results by Merton (1969) and Samuelson (1969) obtain in special cases such as power utility with constant investment opportunities. Alternative solution methods to (1) under predictability of returns are described in Ang and Bekaert (2002), Barberis (2000), Brandt (1999), Brennan, Schwarz and Lagnado (1997), Campbell and Viceira (1999, 2001), Campbell, Chan and Viceira (2003), Kandel and Stambaugh (1996), and Lynch (2001).

[^4]:    ${ }^{7}$ The notation $\kappa_{n, T}$ makes it explicit that the coefficients of the fourth order Taylor expansion depend on the investment horizon through the coefficient $v_{T}$, the point around which the approximation is calculated. We follow standard practice (e.g. Jondeau and Rockinger (2004)) and set the point around which the Taylor series expansion is computed to $v_{T}=E_{t}\left[W_{t+T-1}\right]$, the expected value of the investor's wealth for a $T-1$ period investment horizon.

[^5]:    ${ }^{8}$ Diversification benefits for emerging market investments are highlighted in earlier literature (De Santis (1993), Harvey (1995)) that also discusses how integration leads to increased correlations (Bekaert and Harvey (1997)).

[^6]:    ${ }^{9} K_{i, i, i,-i}$ measures the signed linear association between cubic and simple deviations from means for a pair of assets. A security $i$ with positive values of $K_{i, i, i,-i}$ becomes skewed to the left when other securities pay below-normal returns and is hence undesirable to risk-averse investors.
    ${ }^{10}$ These test results differ from those in Guidolin and Timmermann (2008) (aob? ), who find evidence of non normality also for UK and the Jarque-Bera statistics for Pacific ex-Japan and US are higher than in our dataset- namely 5655.6 vs. 40.94 for Pacific and 162.71 vs. 13.309 for US.

[^7]:    ${ }^{11}$ Catao and Timmermann (2007) construct pure country and pure industry factor mimicking portfolios out of firm level data. By comparison, they reject both linearity and normality in both country and industry returns. A two-regime specification is the most suitable according to three information criteria (BIC, AIC, HQ).
    ${ }^{12}$ This evidence is consistent with previous findings. For instance, Schwert (1989) and Hamilton and Lin (1996) indicate that the volatility of stock returns is higher during recessions than during expansions.
    ${ }^{13}$ This is broadly consistent with Catao and Timmermann (2007), where returns stay for 40 months in bear states and 42 in the normal one for country indexes vs 8 and 26 for industry indexes.

[^8]:    ${ }^{14}$ This last observation does not align with previous findings (on weekly data) showing that defensive US industries have lower correlations in bear markets (Ang and Chen, 2002).
    ${ }^{15}$ The risk premium on value firm may be related to irreversibility of fixed investment. They may be stuck with unproductive capital in bad states of the world and hence underperform with respect to growth firms when the price of risk is high (Zhang, 2005).

[^9]:    ${ }^{16}$ Both MV(2) and MVS investors hold such a three-country portfolio provided their horizon is not lower than 3 months.

[^10]:    ${ }^{17}$ In our comments below we do not emphasize results concerning the 10 year horizon. This is because we produce only 7 (36) out of sample performances for the International (Book-to-Market) experiment. We are in the process of producing results for $\mathrm{T}=60$ horizon, which will deliver a more meaningful comparison.
    ${ }^{18}$ We follow Ang and Bekaert (2002) and Guidolin and Timmerman (2008a) to obtain estimates of the CEQ.
    ${ }^{19}$ This is the idea behind he manipulation of Sharpe ratios discussed by Goetzmann et al.

[^11]:    ${ }^{20}$ We do not actually report performance for $\mathrm{T}=120$ in the International and Emerging dataset, as they get nonsensically large due to the few out-of-sample observations.
    ${ }^{21}$ This pattern is not confirmed only when $T=120$, as MV(1) outperforms both MVS and MV(2). Yet the very few (7) out-of-sample observations in this specific experiment make this evidence unreliable.
    ${ }^{22}$ Such good performance of a MVSK strategy is not necessarily associated wih better ex-post higher order moments: the kurtosis of wealth is actually higher for MVSK than, say, for MV(2) for both $\mathrm{T}=1$ and $\mathrm{T}=12$.

[^12]:    ${ }^{23}$ An exception concerns the equally weighted strategy. While ranking second to MVS - according to Sortino Ratio when $\mathrm{T}=60$ and to both Sharpe and Sortino Ratios for $\mathrm{T}=120$, the confidence bounds of $1 / \mathrm{N}$ and MVS overlap.
    ${ }^{24}$ MVSK is dominated by the equally weighted strategy for $T=120$ according to all the performance measures - the Sharpe Ratios being, for instance, equal to [2.692,5.323] and [0.848,1.306] respectively.

[^13]:    ${ }^{25}$ Both the Treynor ratio and Jensen's alpha show lower ranking stability than the other performance measures. Despite this, Table 7 reveals the presence of some regularities for longer time horizons ( $\mathrm{T} \geq 12$ months).

